

Online Appendix for “Bayesian Compressed Vector Autoregressions”

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Outline

This not-for-publication appendix presents a number of additional results and robustness analysis that were not included in the main paper. The structure of the document is as follows.

[Appendix A](#) presents additional results to highlight the predictive accuracy of the various models considered in the paper, focusing one-by-one on the key seven series of interest, i.e. PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, and GS10. Results on point forecast accuracy are presented in [Figure A.1](#) to [Figure A.6](#), while results for density forecast accuracy are in [Figure A.7](#) to [Figure A.12](#).

[Appendix B](#) presents results for an Intermediate VAR with 46 variables. [Table B.1](#) and the left side of [Table B.3](#) present evidence on the quality of point forecasts for our seven main variables of interest relative to the AR(1) benchmark. [Figure B.1](#) presents evidence on when the forecasting gains of the Intermediate BCVARs are achieved. [Table B.2](#) and the right hand side of [Table B.3](#) shed light on the quality of the density forecasts of the Intermediate VAR by presenting averages of log predictive likelihoods. [Figure B.2](#) plots the cumulative sums of the multivariate log predictive likelihood differentials for the Intermediate VAR across a number of forecast horizons. Finally, [Table B.4](#) shows the forecast performance of our BCVAR_{tvp-sv} approach in the Intermediate VAR case.

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[Appendix C](#) presents results for a number of additional analysis and robustness checks we considered. [Figure C.1](#) and [Figure C.2](#) plot the empirical distributions of m and φ , which define the dimension and the degree of sparsity in the compression matrix, for the Bayesian compressed VARs of different dimensions. They can be interpreted as approximations to the posteriors of these parameters.¹ Next, [Figure C.3](#) and [Figure C.4](#) compare the estimated coefficients of the medium BCVAR using our triangularization scheme for the VAR (defined in equations (6) to (8) of the paper) and the triangularization scheme proposed by Carriero, Clark and Marcellino (2016b), while [Table C.1](#) provides a forecast comparison of the two triangularization schemes. In the table, we label with BCVAR_{ccm} the BCVAR model estimated using the Carriero, Clark, Marcellino (2016b) approach without compressing the covariance terms. Similarly, we label with $\text{BCVAR}_{ccm,c}$ the version of the model where we also compress the covariance terms.² Next, [Figure C.5](#) compares the impulse responses functions from the BCVAR approach to those obtained using OLS, in the case of the medium VAR. [Table C.2](#) presents results for the out-of-sample forecast performance of the BCVAR methods when considering an alternative weighting scheme to perform BMA, where for each equation of the VAR the forecasts resulting from the different random compressions are averaged according to the univariate BIC computed separately for that equation. In the table, we indicate with BCVAR_{alt} and $\text{BCVAR}_{c,alt}$ the two versions of BCVAR, with or without compression on the covariance terms, of this alternative BMA scheme.

To shed light on whether there are statistically significant differences between the multivariate approaches, [Table C.3](#) shows the forecast performance using the BVAR as the benchmark. [Table C.4](#) (which is of the same format and should be compared to [Table B.3](#)) summarizes the forecasting performance of our BCVAR approach when the BICs are calculated using the likelihood of the 7 key variables of interest. [Table C.5](#) presents results with the 7 variables of interest ordered last (labeled $\text{BCVAR}_{c,v,2}$) compared to those of other

¹To aid in interpretation note that, in our compressed VARs, there is a different compression matrix in each equation and so, for brevity, the figures average over all equations and are based on the 75% of draws with highest posterior probability. Remember that smaller values of m indicate a higher degree of compression and, for a given m , $\varphi = 0.5$ induces the highest degree of sparsity.

²The triangularization proposed by Carriero, Clark and Marcellino (2016b) rewrites the original VAR(1) in equation (4) of the paper as $Y_t = BY_{t-1} + \hat{A}^{-1}\Sigma(\tilde{E}_t) + \Sigma E_t$, where, using the notation defined in the paper, $\hat{A}^{-1} = A^{-1} - I_n$, and \tilde{E}_t are residuals which can be estimated recursively (because \hat{A}^{-1} is lower triangular). Compared to our proposed triangularization, defined in equations (6) to (8) of the paper, this form does not multiply B with A^{-1} , rather it retains the original matrix of VAR coefficients.

approaches. [Table C.6](#), [Table C.7](#), and [Table C.8](#) repeat these robustness checks for TVP-SV versions of our approach. [Table C.9](#) presents forecasting results for the BCVAR model with time variation in both the coefficient and covariance matrix (labeled BCVAR_{tvp-sv} in the table) as well as those with only variation in the error covariance matrix (labeled BCVAR_{sv} in the table). Finally, the top two panels of [Table C.9](#) compare our BCVAR_{tvp-sv} and BCVAR_{sv} approaches to the model of Carriero, Clark and Marcellino (2016b) (labeled BVAR_{ccm}).³

Lastly, [Appendix D](#) lists the variable definitions and transformation codes for the 129 variables for which complete data was available.

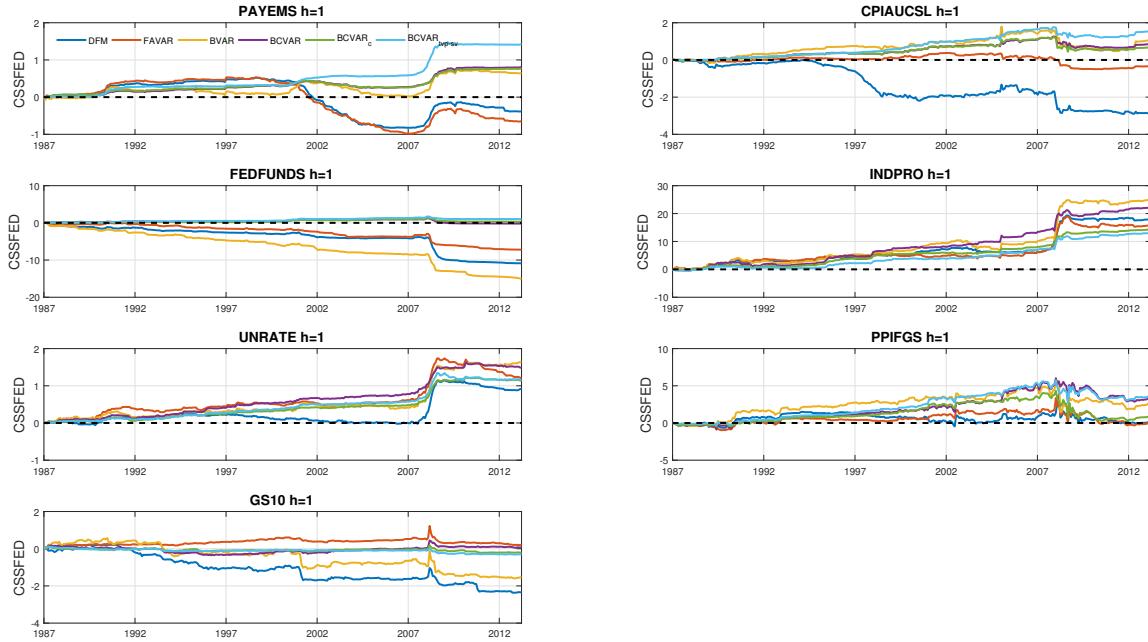
References

- Carriero, A., Clark, T. and Marcellino, M. (2016b). Large vector autoregressions with stochastic volatility and flexible priors, Federal Reserve Bank of Cleveland Working Paper, no. 16-17.

³For the autoregressive coefficients we use the asymmetric Minnesota prior with shrinkage hyperparameter $\lambda = 0.01$, and prior mean for own lags $\delta = 0.95$. For all other parameters our priors are fairly non-informative and are exactly the same as in Carriero, Clark and Marcellino (2016b). Also, note that the results in [Table C.9](#) for the Medium VAR took 25 hours to run on a PC using a modern Core i7 and 32Gb of RAM.

Appendix A Predictive performance for individual series

Figure A.1. Cumulative sum of squared forecast error differentials, Medium VAR, $h = 1$

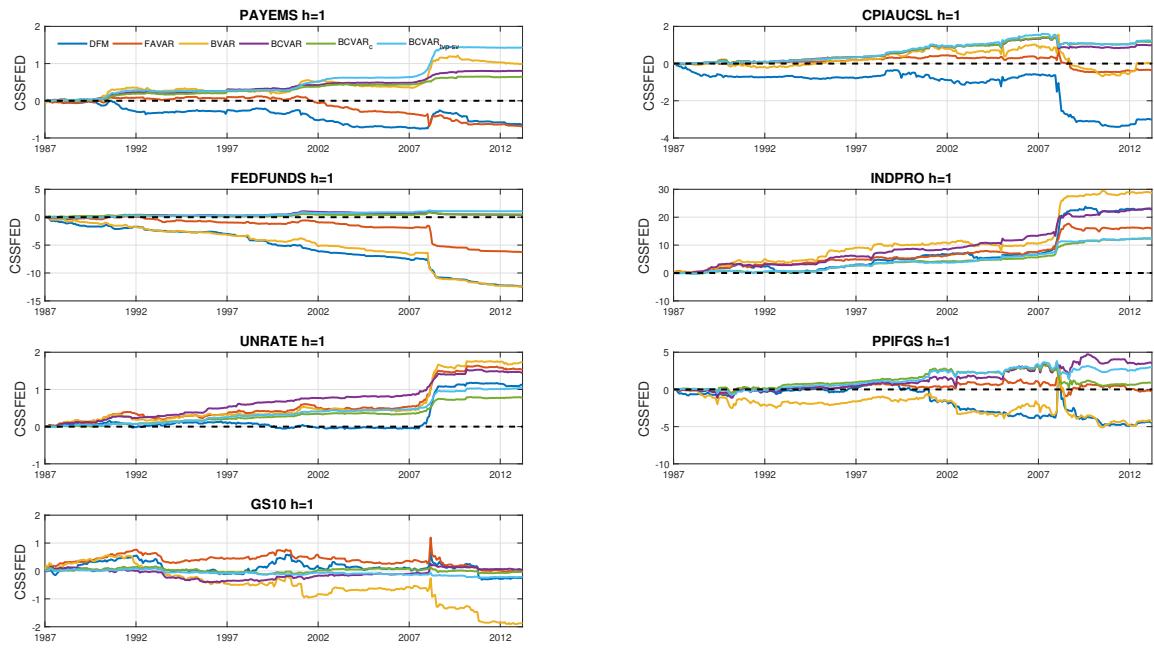


This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for a Medium size VAR and forecast horizon $h = 1$,

$$CSSFED_{ijht} = \sum_{\tau=t}^t (e_{bcmk,j,\tau+h}^2 - e_{i,j,\tau+h}^2)$$

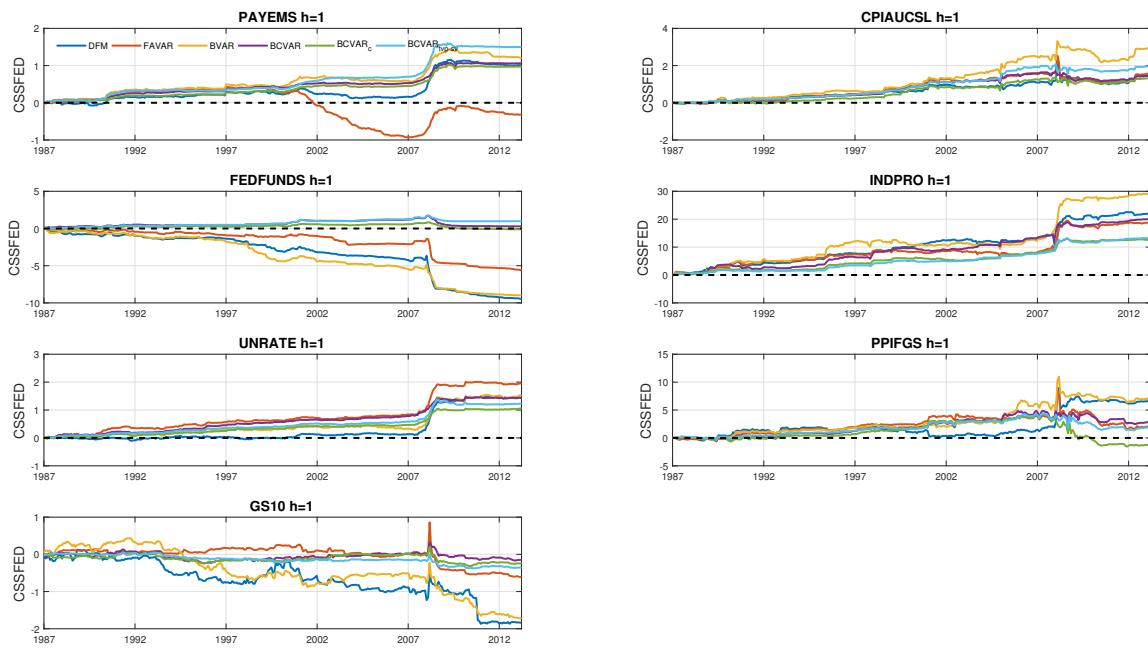
where $t = \underline{t}, \dots, \bar{t}-h$. Values above zero indicate that model i generates better performance than the benchmark, while negative values suggest the opposite. $i \in \{\text{DFM, FAVAR, BVAR, BCVAR, BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$, $j \in \{\text{PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}\}$, \underline{t} and \bar{t} denote the start and end of the out-of-sample period. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Each panel displays results for a different series.

Figure A.2. Cumulative sum of squared forecast error differentials, Intermediate VAR, $h = 1$



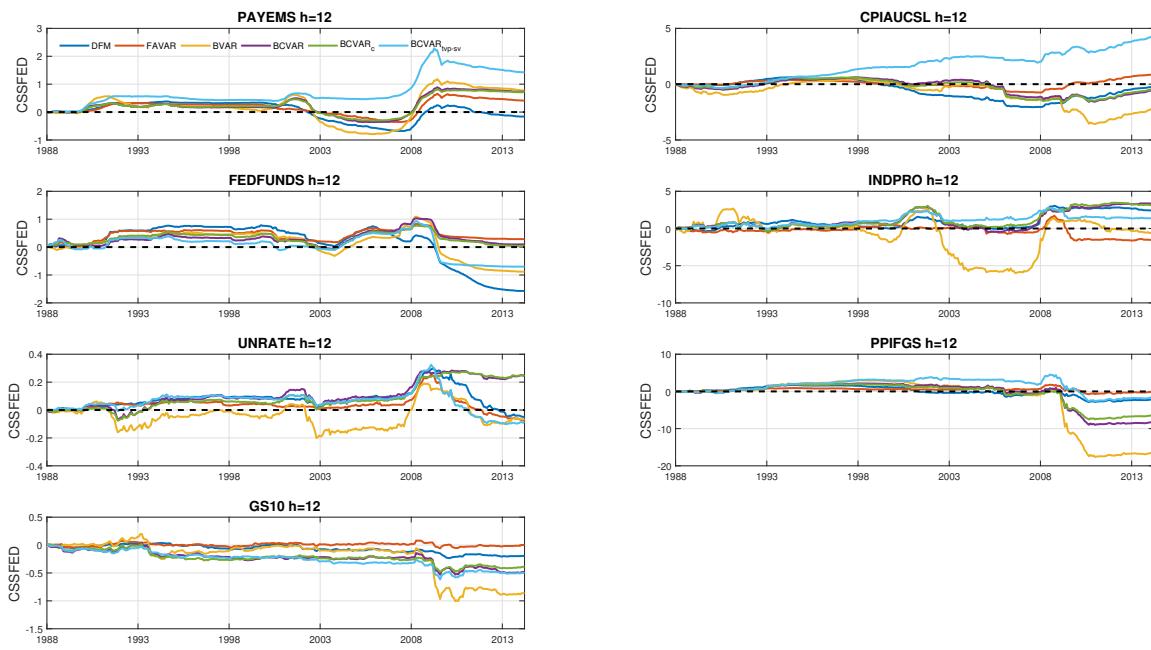
This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for an Intermediate size VAR and forecast horizon $h = 1$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to [Figure A.1](#) for additional details.

Figure A.3. Cumulative sum of squared forecast error differentials, Large VAR, $h = 1$



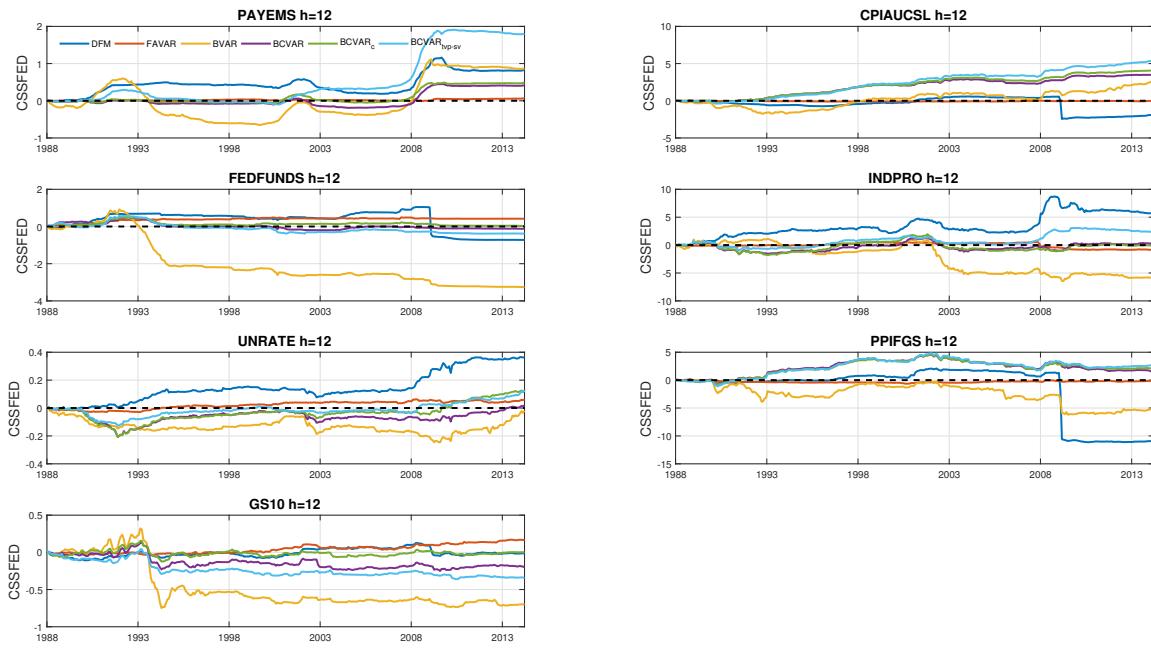
This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for a Large size VAR and forecast horizon $h = 1$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.1 for additional details.

Figure A.4. Cumulative sum of squared forecast error differentials, Medium VAR, $h = 12$



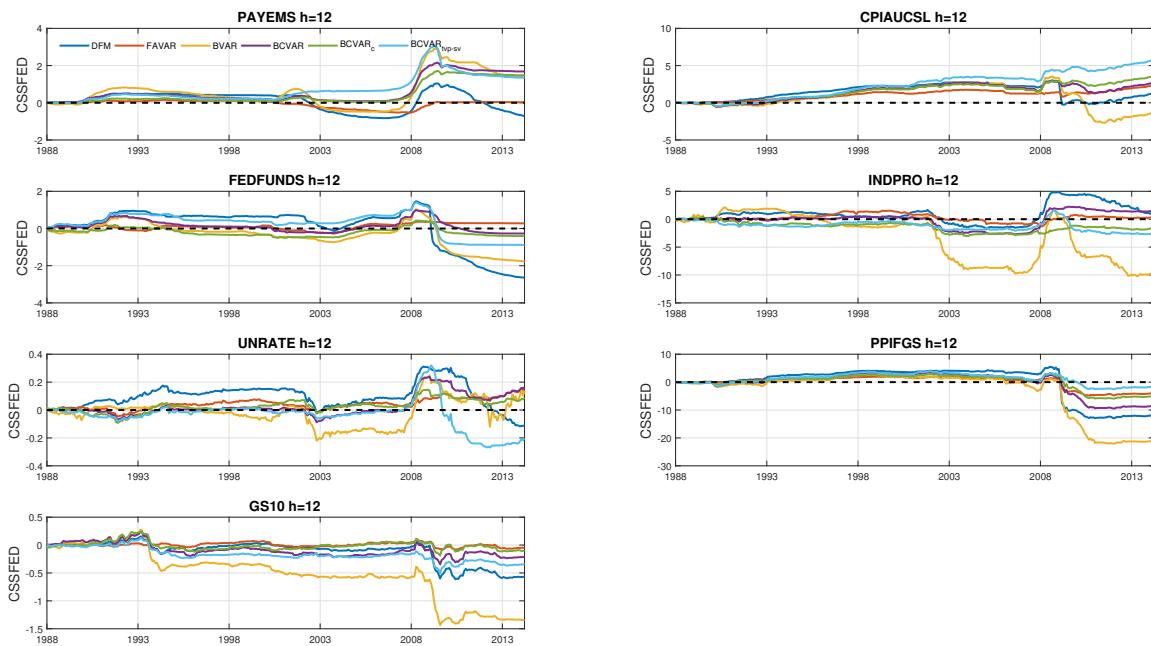
This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for a Medium size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to [Figure A.1](#) for additional details.

Figure A.5. Cumulative sum of squared forecast error differentials, Intermediate VAR, $h = 12$



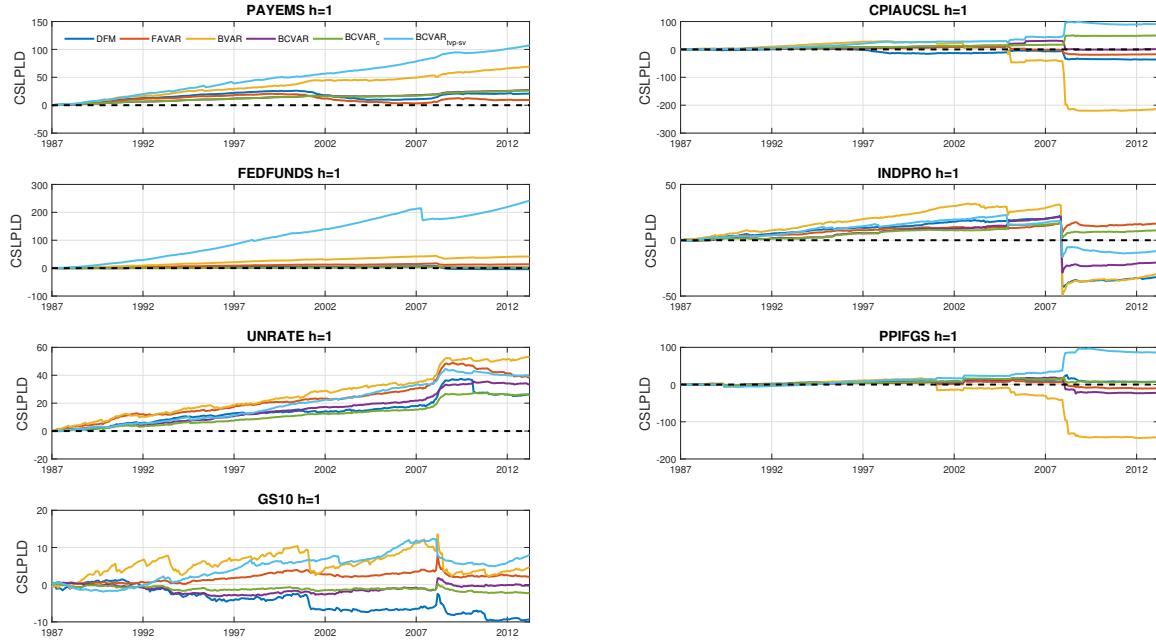
This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for an Intermediate size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to [Figure A.1](#) for additional details.

Figure A.6. Cumulative sum of squared forecast error differentials, Large VAR, $h = 12$



This figure plots the cumulative sum of squared forecast errors generated by the AR(1) model minus the cumulative sum of squared forecast errors generated by model i for a Large size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.1 for additional details.

Figure A.7. Cumulative sum of log predictive likelihood differentials, Medium VAR, $h = 1$

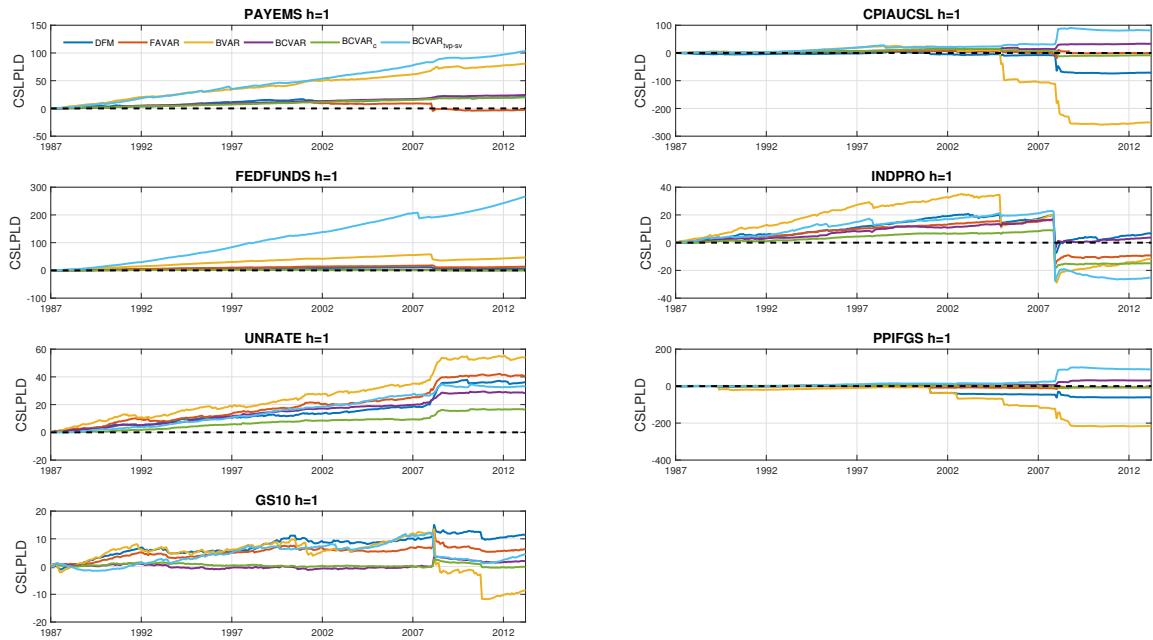


This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for a Medium size VAR and forecast horizon $h = 1$,

$$CSLPLD_{ijht} = \sum_{\tau=t}^t (LPL_{i,j,\tau+h} - LPL_{bcmk,j,\tau+h})$$

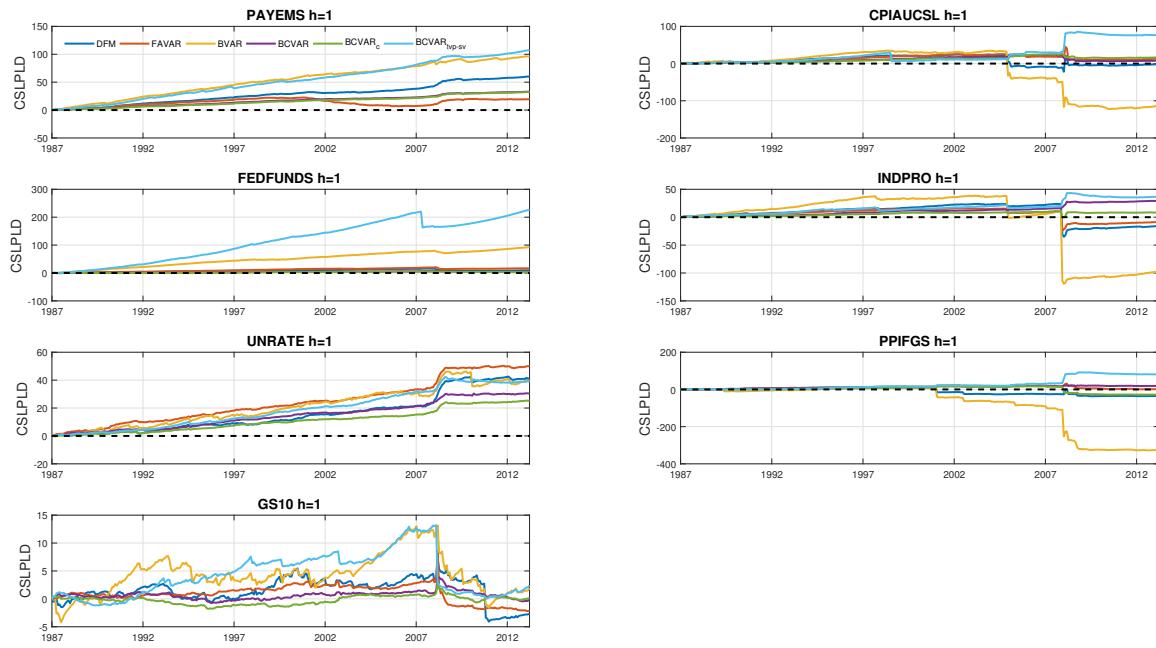
where $t = \underline{t}, \dots, \bar{t} - h$. Values above zero indicate that model i generates better performance than the benchmark, while negative values suggest the opposite. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$, $j \in \{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$, \underline{t} and \bar{t} denote the start and end of the out-of-sample period. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Each panel displays results for a different series.

Figure A.8. Cumulative sum of log predictive likelihood differentials, Intermediate VAR, $h = 1$



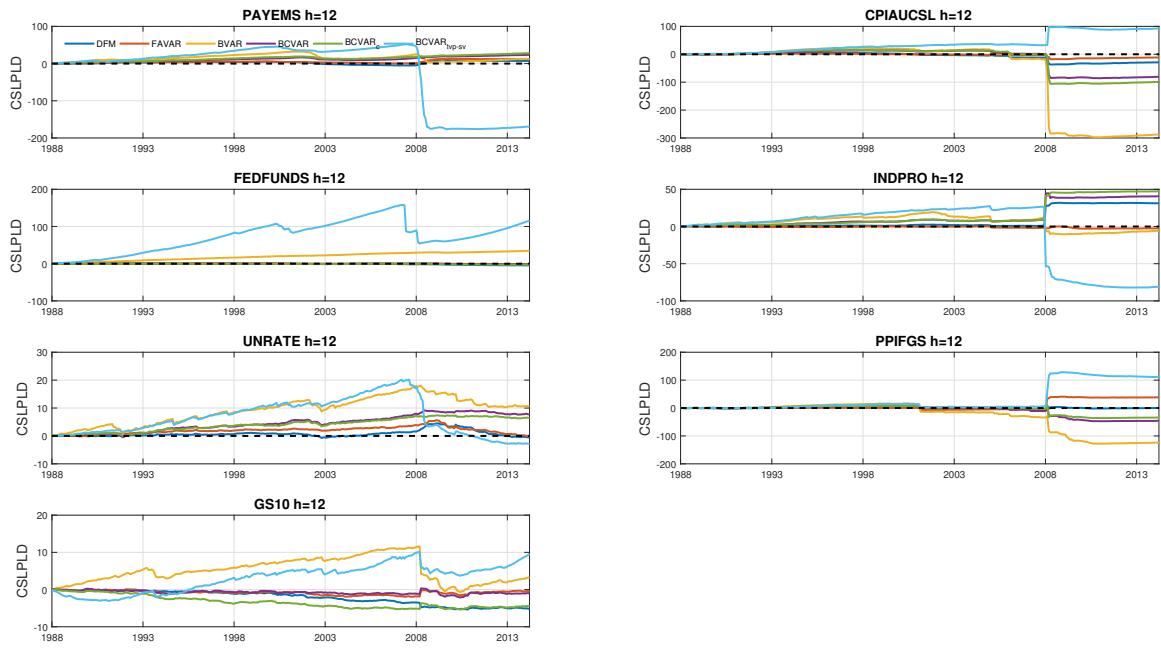
This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for an Intermediate size VAR and forecast horizon $h = 1$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.7 for additional details.

Figure A.9. Cumulative sum of log predictive likelihood differentials, Large VAR, $h = 1$



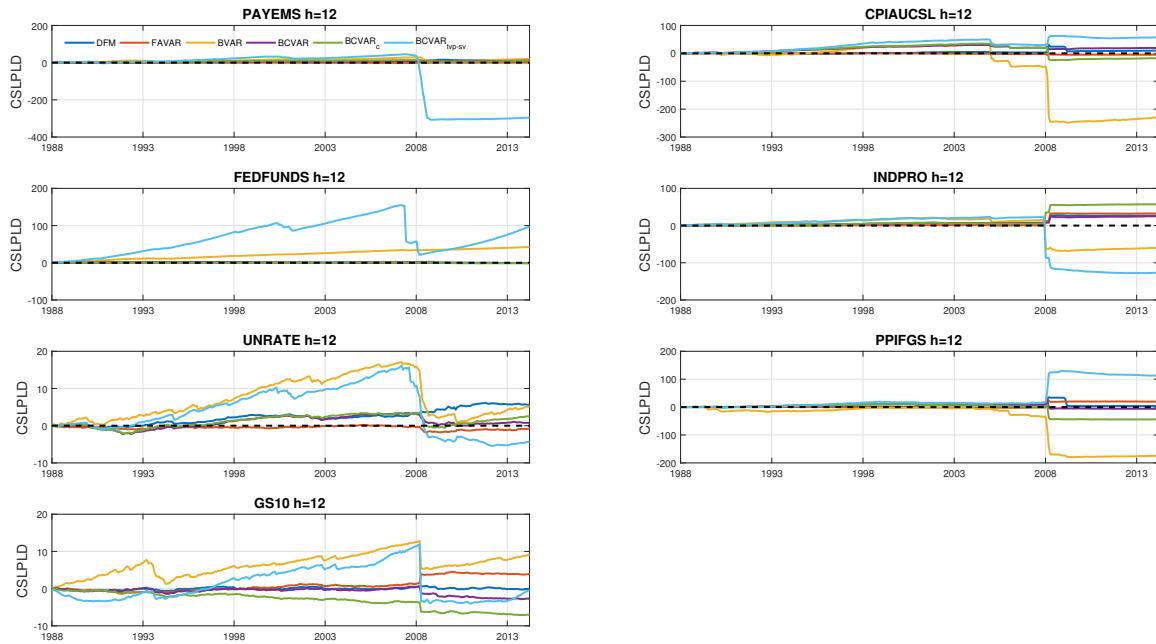
This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for a Large size VAR and forecast horizon $h = 1$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.7 for additional details.

Figure A.10. Cumulative sum of log predictive likelihood differentials, Medium VAR, $h = 12$



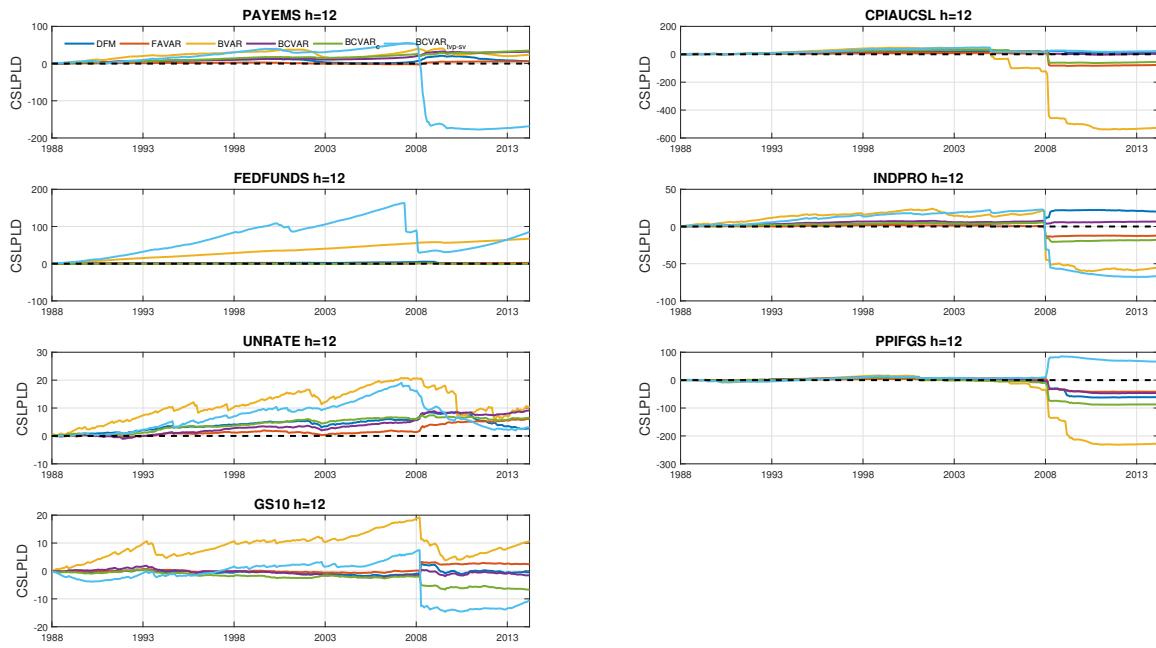
This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for a Medium size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.7 for additional details.

Figure A.11. Cumulative sum of log predictive likelihood differentials, Intermediate VAR, $h = 12$



This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for an Intermediate size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.7 for additional details.

Figure A.12. Cumulative sum of log predictive likelihood differentials, Large VAR, $h = 12$



This figure plots the cumulative sum of log predictive likelihoods generated by model i minus the cumulative sum of log predictive likelihoods generated by the AR(1) model for a Large size VAR and forecast horizon $h = 12$. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c, \text{BCVAR}_{tvp-sv}\}$. See notes to Figure A.7 for additional details.

Appendix B Results for Intermediate VAR

Table B.1. Out-of-sample point forecast performance, Intermediate VAR

| Variable | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | |
|----------|----------------|--------------|-----------------|----------------|--------------------|----------------|---------------|-----------------|--------------|--------------------|--|
| | <i>h</i> = 1 | | | | | | <i>h</i> = 2 | | | | |
| PAYEMS | 1.137 | 1.146 | 0.792** | 0.831*** | 0.864*** | 0.869 | 0.914 | 0.512*** | 0.747*** | 0.762*** | |
| CPIAUCSL | 1.148 | 1.017 | 1.000 | 0.951 | 0.942* | 1.165 | 1.085 | 1.099 | 0.911** | 0.898*** | |
| FEDFUNDS | 2.449 | 1.731 | 2.449 | 0.949 | 0.944 | 1.961 | 1.376 | 2.532 | 0.963 | 0.924 | |
| INDPRO | 0.824** | 0.877*** | 0.778*** | 0.820*** | 0.904*** | 0.855 | 0.918* | 0.771** | 0.907** | 0.935* | |
| UNRATE | 0.851** | 0.798*** | 0.770*** | 0.809*** | 0.897*** | 0.803** | 0.841*** | 0.794** | 0.857*** | 0.893*** | |
| PPIFGS | 1.042 | 1.002 | 1.041 | 0.967 | 0.991 | 1.157 | 1.057 | 1.166 | 1.013 | 1.006 | |
| GS10 | 1.015 | 1.001 | 1.113 | 0.997 | 1.002 | 0.999 | 1.023 | 1.116 | 0.996 | 1.009 | |
| | <i>h</i> = 3 | | | | | | <i>h</i> = 6 | | | | |
| PAYEMS | 0.780 | 0.842* | 0.467*** | 0.717*** | 0.732*** | 0.841 | 0.920* | 0.604** | 0.764** | 0.783*** | |
| CPIAUCSL | 1.132 | 1.061 | 1.146 | 0.923** | 0.926** | 1.045 | 1.018 | 0.988 | 0.897*** | 0.885*** | |
| FEDFUNDS | 1.714 | 1.063 | 2.174 | 1.001 | 0.989 | 1.247 | 0.974 | 1.234 | 0.998 | 0.963 | |
| INDPRO | 0.900 | 0.944 | 0.852 | 0.927** | 0.938* | 0.939 | 0.981 | 0.980 | 0.975 | 0.971 | |
| UNRATE | 0.855* | 0.911** | 0.840* | 0.906** | 0.930** | 0.906** | 0.956*** | 0.887** | 0.927** | 0.962 | |
| PPIFGS | 1.143 | 1.003 | 1.168 | 1.004 | 1.007 | 1.104 | 1.008 | 1.088 | 1.001 | 0.993 | |
| GS10 | 1.040 | 1.024 | 1.211 | 1.050 | 1.047 | 1.038 | 1.009 | 1.098 | 1.031 | 1.022 | |
| | <i>h</i> = 9 | | | | | | <i>h</i> = 12 | | | | |
| PAYEMS | 0.877 | 0.962** | 0.762 | 0.858* | 0.863** | 0.926 | 0.994 | 0.922 | 0.962 | 0.956 | |
| CPIAUCSL | 1.047 | 0.998 | 0.910 | 0.848*** | 0.841*** | 1.065 | 1.002 | 0.898 | 0.880*** | 0.860*** | |
| FEDFUNDS | 1.113 | 1.008 | 1.179 | 0.970 | 1.025 | 1.062 | 0.964* | 1.281 | 1.010 | 0.997 | |
| INDPRO | 0.962 | 1.009 | 1.003 | 0.987 | 0.988 | 0.957 | 1.006 | 1.043 | 0.998 | 1.000 | |
| UNRATE | 0.949** | 0.987 | 0.965 | 0.979 | 0.987 | 0.954** | 0.992 | 1.002 | 0.998 | 0.985 | |
| PPIFGS | 1.059 | 1.002 | 1.049 | 0.973 | 0.973 | 1.096 | 1.002 | 1.042 | 0.989 | 0.981 | |
| GS10 | 0.998 | 0.998 | 1.043 | 0.995 | 1.022 | 1.001 | 0.990 | 1.043 | 1.012 | 1.000 | |

This table reports the ratio between the MSFE of model *i* and the MSFE of the benchmark AR(1) for the Intermediate VAR, computed as

$$MSFE_{ijh} = \frac{\sum_{\tau=t}^{\bar{t}-h} e_{i,j,\tau+h}^2}{\sum_{\tau=t}^{\bar{t}-h} e_{bcmk,j,\tau+h}^2},$$

where $e_{i,j,\tau+h}^2$ and $e_{bcmk,j,\tau+h}^2$ are the squared forecast errors of variable *j* at time τ and forecast horizon *h* generated by model *i* and the AR(1) model, respectively. t and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM, FAVAR, BVAR, BCVAR, BCVAR}_c\}$, $j \in \{\text{PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest MSFE across all models for a given variable-forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table B.2. Out-of-sample density forecast performance, Intermediate VAR

| Variable | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
|----------|---------------|---------------|-----------------|----------------|--------------------|----------------|--------------|-----------------|----------------|--------------------|
| | | | | | | | | | | |
| | | | h = 1 | | | | | h = 2 | | |
| PAYEMS | 0.065*** | -0.008 | 0.254*** | 0.076*** | 0.063*** | 0.138*** | 0.079*** | 0.406*** | 0.140*** | 0.137*** |
| CPIAUCSL | -0.223 | -0.002 | -0.787 | 0.104** | -0.026 | -0.800 | 0.061 | -2.100 | 0.032 | 0.182** |
| FEDFUNDS | 0.022 | 0.042** | 0.147** | 0.004 | -0.002 | -0.002 | 0.018 | -0.026 | 0.000 | 0.009 |
| INDPRO | 0.020 | -0.030 | -0.039 | 0.011 | -0.047 | 0.162*** | 0.064 | 0.181** | 0.065*** | 0.090** |
| UNRATE | 0.114*** | 0.127*** | 0.170*** | 0.089*** | 0.051*** | 0.118*** | 0.080*** | 0.150*** | 0.083*** | 0.058*** |
| PPIFGS | -0.191 | -0.021 | -0.679 | 0.096* | -0.024 | -0.477 | -0.091 | -1.105 | 0.053 | -0.048 |
| GS10 | 0.036* | 0.020 | -0.027 | 0.006 | 0.000 | 0.017 | 0.014 | 0.017 | 0.021 | -0.003 |
| | | | h = 3 | | | | | h = 6 | | |
| PAYEMS | 0.141*** | 0.060*** | 0.416*** | 0.159*** | 0.160*** | 0.082** | 0.029 | 0.300*** | 0.121*** | 0.140*** |
| CPIAUCSL | -0.246 | -0.086 | -1.873 | -0.073 | -0.075 | -0.091 | 0.078 | -0.826 | 0.087** | -0.061 |
| FEDFUNDS | -0.046 | 0.011 | 0.029 | 0.004 | 0.006 | -0.007 | -0.002 | 0.159*** | 0.003 | 0.012* |
| INDPRO | -0.022 | 0.005 | -0.056 | 0.010 | -0.021 | 0.069** | -0.128 | -0.315 | -0.063 | -0.149 |
| UNRATE | 0.081*** | 0.035** | 0.119*** | 0.061*** | 0.054*** | 0.039*** | 0.017** | 0.092*** | 0.042*** | 0.023** |
| PPIFGS | -0.193 | -0.061 | -1.087 | 0.029 | -0.125 | -0.064 | 0.049 | -0.791 | 0.007 | -0.099 |
| GS10 | -0.001 | -0.010 | -0.028 | 0.000 | -0.008 | -0.005 | -0.005 | 0.004 | -0.004 | -0.009 |
| | | | h = 9 | | | | | h = 12 | | |
| PAYEMS | 0.059** | 0.019 | 0.165*** | 0.095*** | 0.097*** | 0.044 | 0.003 | 0.063 | 0.034* | 0.032 |
| CPIAUCSL | -0.157 | -0.040 | -0.872 | -0.104 | -0.158 | 0.032 | -0.016 | -0.721 | 0.059 | -0.058 |
| FEDFUNDS | -0.006 | -0.006 | 0.145*** | -0.002 | -0.003 | -0.001 | -0.002 | 0.133*** | -0.003 | -0.004 |
| INDPRO | 0.085 | 0.029 | -0.178 | 0.027*** | 0.050 | 0.083 | 0.102 | -0.188 | 0.078 | 0.180 |
| UNRATE | 0.033*** | -0.002 | 0.050** | 0.017** | 0.013* | 0.018** | -0.002 | 0.017 | 0.002 | 0.008 |
| PPIFGS | -0.036 | -0.021 | -0.647 | -0.047 | -0.047 | 0.014 | 0.061 | -0.549 | -0.021 | -0.138 |
| GS10 | 0.004 | -0.001 | 0.034 | 0.001 | -0.017 | 0.000 | 0.012 | 0.029 | -0.009 | -0.022 |

This table reports the average log predictive likelihood (ALPL) differential between model i and the benchmark AR(1) for the Intermediate VAR , computed as

$$ALPL_{ijh} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (LPL_{i,j,\tau+h} - LPL_{bcmk,j,\tau+h}),$$

where $LPL_{i,j,\tau+h}$ and $LPL_{bcmk,j,\tau+h}$ are the log predictive likelihoods of variable j at time τ and forecast horizon h generated by model i and the AR(1) model, respectively. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c\}$, $j \in \{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. All density forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the highest ALPL across all models for a given variable-forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table B.3. Out-of-sample forecast performance: Multivariate results

| Fest h. | Intermediate VAR | | | | | MVALPL | | | | |
|---------|------------------|---------|-------|-----------------|--------------------|----------|----------|-----------------|-----------------|--------------------|
| | WMSFE | | | | | | | | | |
| | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | 1.160 | 1.048 | 1.103 | 0.906*** | 0.939*** | 0.710*** | 0.820*** | 0.988*** | 0.933*** | 0.253*** |
| h= 2 | 1.117 | 1.033 | 1.148 | 0.919*** | 0.924*** | 0.847*** | 0.844*** | 0.895*** | 1.011*** | 0.360*** |
| h= 3 | 1.083 | 0.981 | 1.126 | 0.934** | 0.939*** | 0.886*** | 0.835*** | 0.945*** | 1.023*** | 0.264*** |
| h= 6 | 1.016 | 0.980** | 0.977 | 0.937** | 0.935** | 0.937*** | 0.828*** | 1.187*** | 1.054*** | 0.276*** |
| h= 9 | 0.999 | 0.994 | 0.979 | 0.939** | 0.951** | 0.935*** | 0.828*** | 1.198*** | 1.043*** | 0.271*** |
| h=12 | 1.009 | 0.993* | 1.026 | 0.975 | 0.965** | 0.886*** | 0.837*** | 1.017*** | 0.956*** | 0.157* |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, $\{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

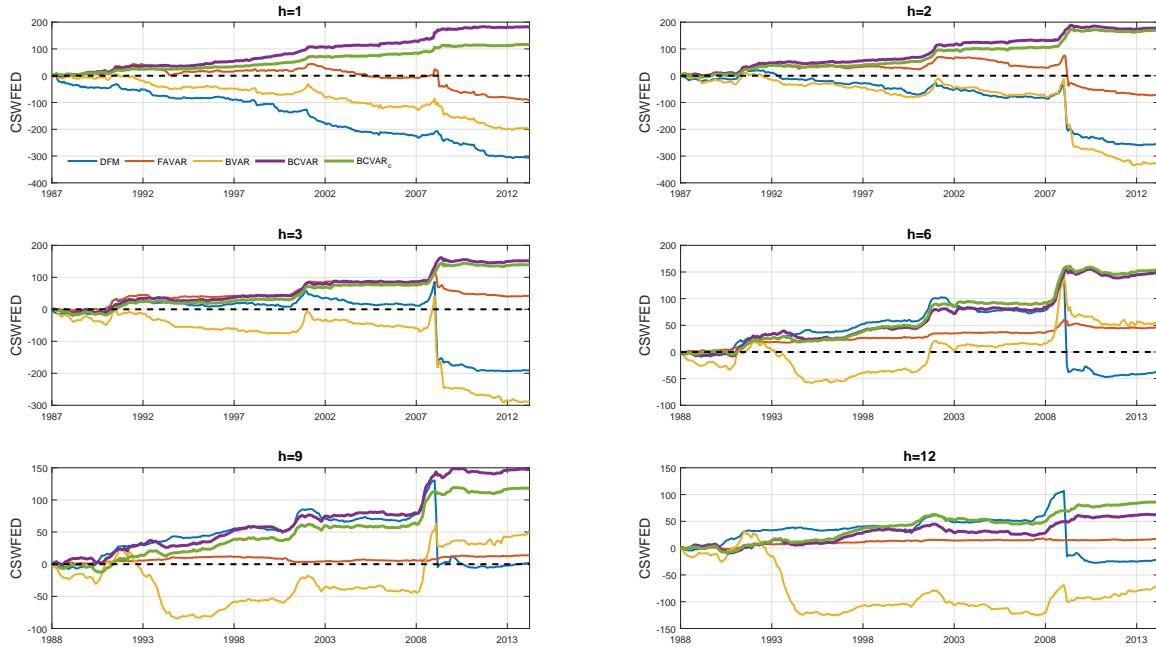
where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given forecast horizon. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table B.4. Out-of-sample forecast performance: Compressed TVP-SV VAR

| Variable | MSFE | | | | | | Intermediate VAR | | | | | ALPL | | | | | | | | | | | | |
|--------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|------------------|-----------------|-----------------|-----------------|---------------|--------------|--------------|--|--------------|--|--------------|--|--------------|--|--------------|--|---------------|--|
| | <i>h</i> = 1 | | <i>h</i> = 2 | | <i>h</i> = 3 | | <i>h</i> = 6 | | <i>h</i> = 9 | | <i>h</i> = 12 | | <i>h</i> = 1 | | <i>h</i> = 2 | | <i>h</i> = 3 | | <i>h</i> = 6 | | <i>h</i> = 9 | | <i>h</i> = 12 | |
| | | | | | | | | | | | | | | | | | | | | | | | | |
| PAYEMS | 0.699*** | 0.566*** | 0.565*** | 0.648** | 0.739** | 0.837 | 0.326*** | 0.387*** | 0.335*** | -0.064 | -0.508 | -0.929 | | | | | | | | | | | | |
| CPIAUCSL | 0.939 | 0.870*** | 0.862*** | 0.843*** | 0.796*** | 0.809*** | 0.257 | 0.486 | 0.306 | 0.244 | 0.345 | 0.181 | | | | | | | | | | | | |
| FEDFUND\$ | 0.875** | 0.847** | 0.843** | 0.932 | 0.968 | 1.033 | 0.838*** | 0.616** | 0.531* | 0.380 | 0.073 | 0.307 | | | | | | | | | | | | |
| INDPRO | 0.904*** | 0.930* | 0.936* | 0.962 | 0.983 | 0.982 | -0.079 | -0.085 | -0.189 | -0.348 | -0.290 | -0.399 | | | | | | | | | | | | |
| UNRATE | 0.862*** | 0.863*** | 0.899** | 0.926** | 0.959** | 0.984 | 0.104*** | 0.104*** | 0.078*** | 0.052*** | 0.031 | -0.013 | | | | | | | | | | | | |
| PPIFGS | 0.972 | 0.985 | 0.983 | 0.987 | 0.958 | 0.976 | 0.285 | 0.400 | 0.371 | 0.361 | 0.379 | 0.359 | | | | | | | | | | | | |
| GS10 | 1.013 | 1.007 | 1.037 | 1.023 | 1.012 | 1.021 | 0.015 | 0.009 | -0.049 | -0.008 | -0.013 | -0.001 | | | | | | | | | | | | |
| Multivariate | 0.910*** | 0.878*** | 0.877*** | 0.896*** | 0.908*** | 0.941** | 1.633*** | 1.635*** | 1.511*** | 1.215*** | 0.966*** | 0.674 | | | | | | | | | | | | |

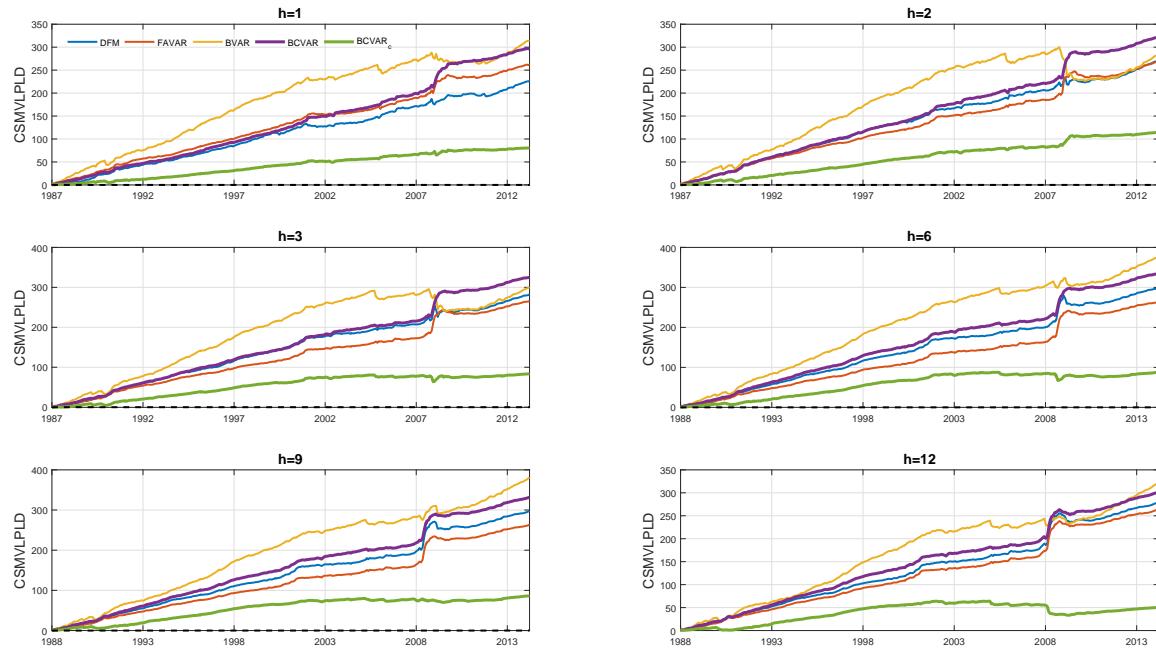
The left half of this table reports the ratio between the univariate or multivariate weighted mean squared forecast error of the BCVAR_{*tvp-sv*} model and the univariate or multivariate weighted mean squared forecast error of the benchmark AR(1) model. The right half of the table shows the univariate or multivariate average log predictive likelihood differentials between the BCVAR_{*tvp-sv*} model and the benchmark AR(1) model. *h* denotes the forecast horizons, with *h* ∈ {1, 2, 3, 6, 9, 12}. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate all instances where the BCVAR_{*tvp-sv*} model outperforms all alternative models (DFM, FAVAR, BVAR, BCVAR, BCVAR_c), for any given variable/forecast horizon combination. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Figure B.1. Cumulative sum of weighted forecast error differentials, Intermediate VAR



This figure plots the cumulative sum of weighted forecast errors generated by the AR(1) model minus the cumulative sum of weighted forecast errors generated by model i for the Intermediate VAR. We define the weighted forecast error of model i and the AR(1) model at time $\tau + h$ as $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$, where $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, $\{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Each panel displays results for a different forecast horizon.

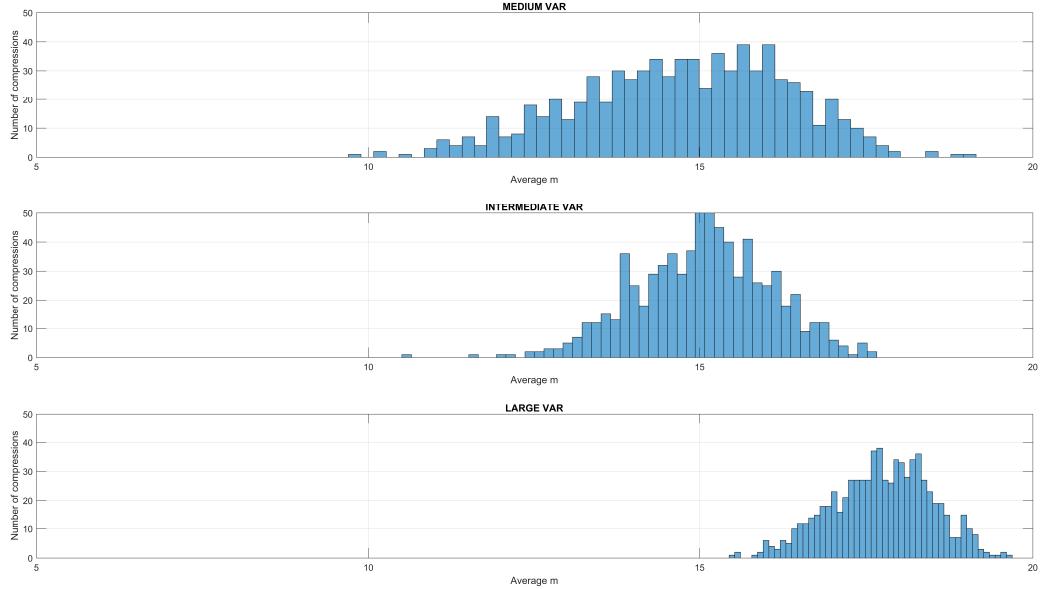
Figure B.2. Cumulative sum of multivariate log predictive likelihood differentials, Intermediate VAR



This figure plots the cumulative sum of the multivariate log predictive likelihoods generated by model i minus the cumulative sum of the multivariate log predictive likelihoods computed from an AR(1) model for the Intermediate VAR. $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c\}$, $h \in \{1, 2, 3, 6, 9, 12\}$, and the multivariate log predictive likelihoods are computed under the assumption of joint normality, as described in the text. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Each panel displays results for a different forecast horizon.

Appendix C Additional analysis

Figure C.1. Average compression size (m) for top 75% compressions

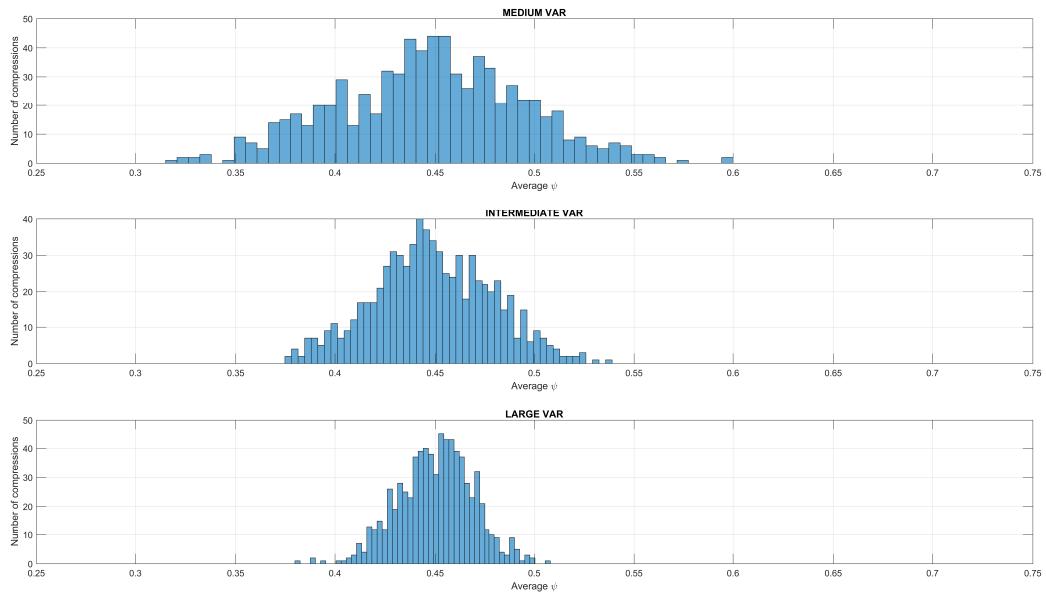


This figure displays the empirical distribution of the average number of rows of the random compression matrices Φ_i , $i = 1, \dots, n$ averaged across all n equations of the VAR, and according to the top 75% compressions, ranked in terms of the VAR overall BIC. For each of the n equations in the VAR, the model specification is

$$Y_{i,t} = \Theta_i^c (\Phi_i Z_t^i) + \sigma_i E_{i,t} \quad i = 1, \dots, n$$

where Z_t^i denotes the subset of the vector Z_t which applies to the i -th equation of the VAR: $Z_t^1 = (Y_{t-1})$, $Z_t^2 = (Y_{t-1}', -Y_{1,t})'$, $Z_t^3 = (Y_{t-1}', -Y_{1,t}, -Y_{2,t})'$, and so on. Similarly, Φ_i is a matrix with m rows and column dimension that conforms with Z_t^i .

Figure C.2. Average sparsity (φ) for top 75% compressions

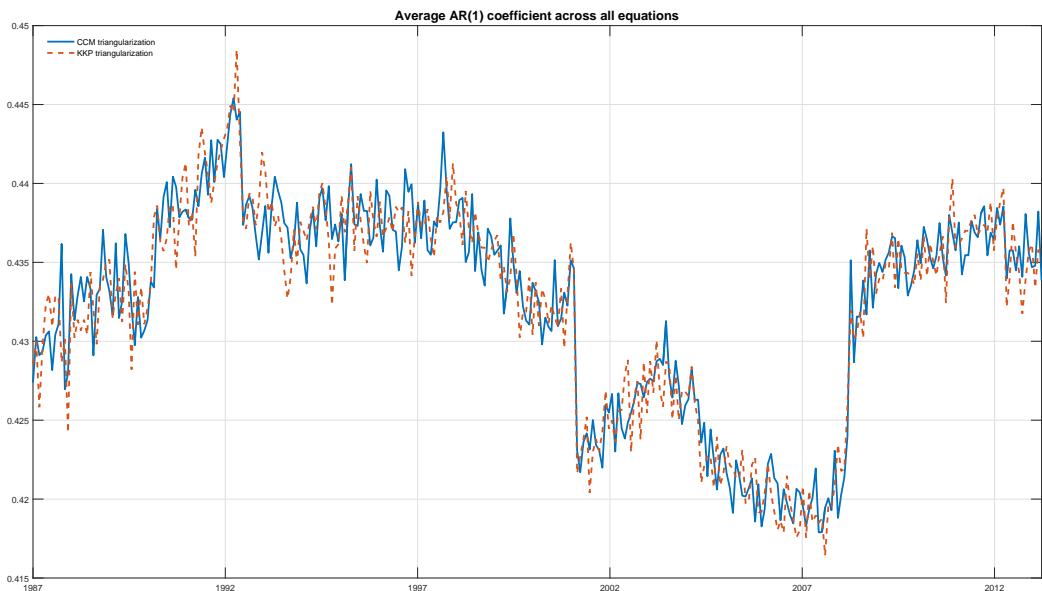


This figure displays the empirical distribution of the average sparsity factor φ of the matrix Φ_i , $i = 1, \dots, n$ averaged across all n equations of the VAR, and according to the top 75% compressions, ranked in terms of the VAR overall BIC. For each of the n equations in the VAR, the model specification is

$$Y_{i,t} = \Theta_i^c (\Phi_i Z_t^i) + \sigma_i E_{i,t} \quad i = 1, \dots, n$$

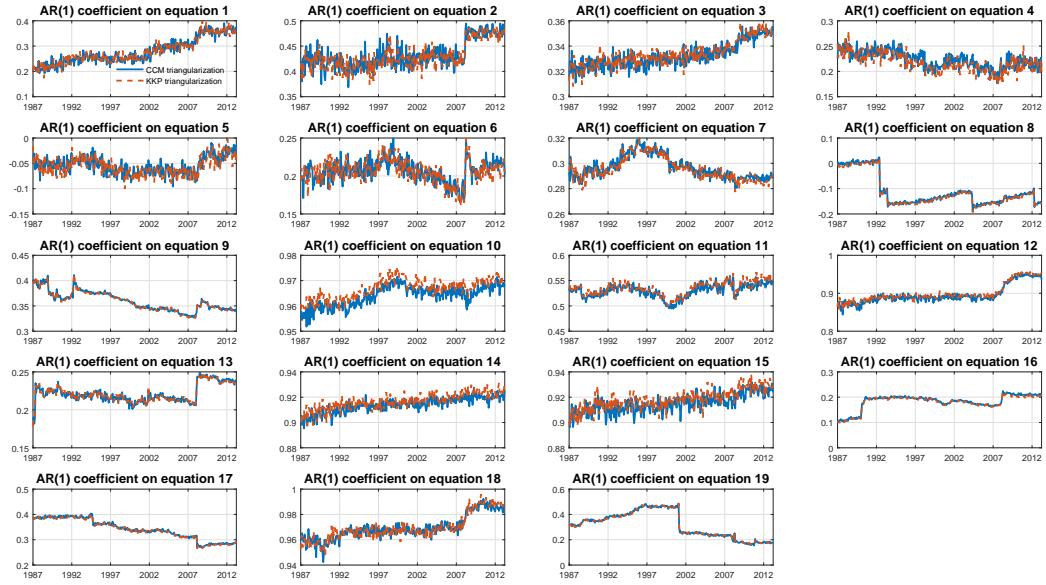
where Z_t^i denotes the subset of the vector Z_t which applies to the i -th equation of the VAR: $Z_t^1 = (Y_{t-1})'$, $Z_t^2 = (Y_{t-1}', -Y_{1,t})'$, $Z_t^3 = (Y_{t-1}', -Y_{1,t}, -Y_{2,t})'$, and so on. Similarly, Φ_i is a matrix with m rows and column dimension that conforms with Z_t^i .

Figure C.3. Average AR(1) coefficients across equations, Medium VAR



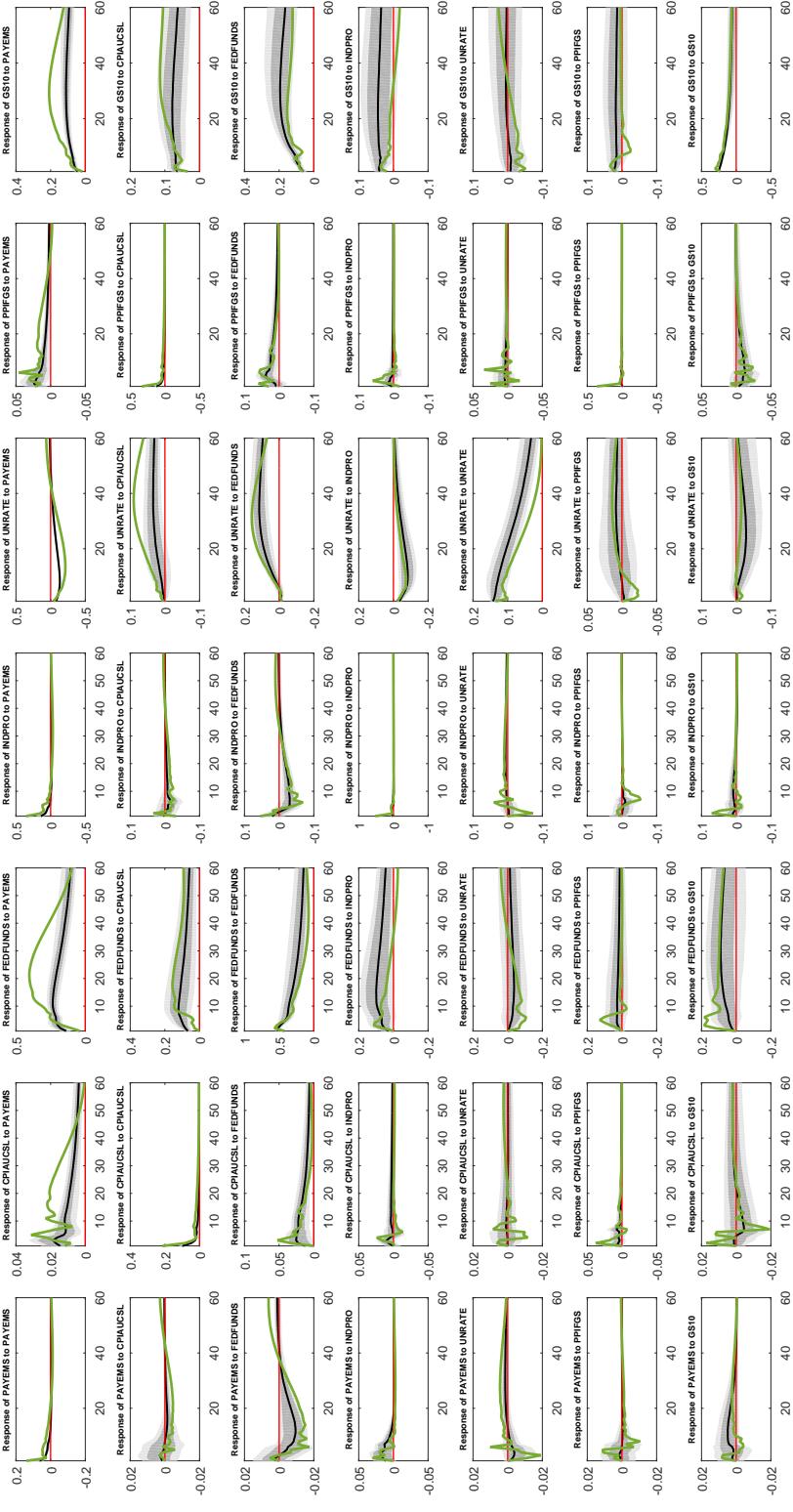
This figure displays the average of all 19 first own lag coefficients using the Carriero, Clark and Marcellino (2016b) and our triangularization algorithms in the Medium VAR. For the first method this is the average of all elements B_{ii} for $i = 1, \dots, 19$ where $B = B^c\Phi$. For the second method this is the average of all elements B_{ii} for $i = 1, \dots, 19$ where $B = A^{-1}(\Gamma^c\Phi)$. The x-axis represents the 318 observations in the evaluation sample.

Figure C.4. AR(1) coefficients across equations, Medium VAR



This figure displays each of the 19 first own lag coefficients using the Carriero, Clark and Marcellino (2016b) and our triangularization algorithms in the Medium VAR. For the first method these are the elements B_{ii} for $i = 1, \dots, 19$ where $B = B^c\Phi$. For the second method these are the elements B_{ii} for $i = 1, \dots, 19$ where $B = A^{-1}(\Gamma^c\Phi)$. The x-axis represents the 318 observations in the evaluation sample.

Figure C.5. Impulse response functions, Medium VAR



This figure depicts the impulse response functions for the seven variables of interest (PAYEMS, CPIAUCSL, UNRATE, PPIFGS, GS10), obtained using the Medium VAR. Shaded areas denote the 68% posterior probability intervals of the impulse response functions from the Medium BCVAR for the seven variables of interest, with the posterior median represented by the black line. Green lines are the mean impulse responses from the same VAR estimated with OLS.

Table C.1. Out-of-sample forecast performance: Multivariate results, alternative triangularization schemes

| Fcst h. | | | | Medium VAR | | | | | | | |
|------------------|---------------|-----------------|-----------------|-----------------|---------------|-----------------|-----------------|-----------|--|--|--|
| WMSFE | | | | MVALPL | | | | | | | |
| | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | | | |
| h= 1 | 0.941*** | 0.916*** | 0.934*** | 0.935*** | 0.659*** | 0.925*** | 0.300*** | 0.285*** | | | |
| h= 2 | 0.933*** | 0.929** | 0.925*** | 0.926*** | 0.806*** | 1.021*** | 0.426*** | 0.401*** | | | |
| h= 3 | 0.949* | 0.944* | 0.950 | 0.940* | 0.810*** | 1.046*** | 0.417*** | 0.356*** | | | |
| h= 6 | 0.953 | 0.961 | 0.964 | 0.954 | 0.753*** | 1.009*** | 0.288*** | 0.296*** | | | |
| h= 9 | 0.958 | 0.957 | 0.961 | 0.960 | 0.738*** | 1.017*** | 0.242*** | 0.254*** | | | |
| h=12 | 1.001 | 0.996 | 0.994 | 0.994 | 0.663*** | 0.886*** | 0.182*** | 0.176*** | | | |
| Intermediate VAR | | | | | | | | | | | |
| WMSFE | | | | MVALPL | | | | | | | |
| | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | | | |
| h= 1 | 0.946*** | 0.906*** | 0.939*** | 0.939*** | 0.730*** | 0.933*** | 0.265*** | 0.253*** | | | |
| h= 2 | 0.940*** | 0.919*** | 0.924*** | 0.924*** | 0.851*** | 1.011*** | 0.333*** | 0.360*** | | | |
| h= 3 | 0.936*** | 0.934** | 0.937** | 0.939*** | 0.850*** | 1.023*** | 0.293*** | 0.264*** | | | |
| h= 6 | 0.937*** | 0.937** | 0.941** | 0.935** | 0.882*** | 1.054*** | 0.278*** | 0.276*** | | | |
| h= 9 | 0.949** | 0.939** | 0.941** | 0.951** | 0.879*** | 1.043*** | 0.261*** | 0.271*** | | | |
| h=12 | 0.975** | 0.975 | 0.971* | 0.965** | 0.825*** | 0.956*** | 0.168** | 0.157* | | | |
| Large VAR | | | | | | | | | | | |
| WMSFE | | | | MVALPL | | | | | | | |
| | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | $BCVAR_{ccm}$ | $BCVAR$ | $BCVAR_{ccm,c}$ | $BCVAR_c$ | | | |
| h= 1 | 0.963** | 0.907*** | 0.935*** | 0.940*** | 0.511*** | 0.996*** | 0.290*** | 0.303*** | | | |
| h= 2 | 0.939** | 0.909*** | 0.917** | 0.908*** | 0.648*** | 1.139*** | 0.371*** | 0.406*** | | | |
| h= 3 | 0.934* | 0.916** | 0.923* | 0.922** | 0.697*** | 1.179*** | 0.319*** | 0.368*** | | | |
| h= 6 | 0.942 | 0.933 | 0.949 | 0.940 | 0.633*** | 1.131*** | 0.301*** | 0.269*** | | | |
| h= 9 | 0.946 | 0.938 | 0.939 | 0.943 | 0.608*** | 1.076*** | 0.253*** | 0.243*** | | | |
| h=12 | 0.968 | 0.969 | 0.963 | 0.968 | 0.544*** | 1.009*** | 0.145 | 0.145 | | | |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, {PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{BCVAR_{ccm}, BCVAR, BCVAR_{ccm,c}, BCVAR_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.2. Out-of-sample forecast performance: Multivariate results, alternative BMA schemes

| Fcst h. | | Medium VAR | | | | | | | |
|------------------|--|----------------------|-----------------|------------------------|--------------------|----------------------|-----------------|------------------------|--------------------|
| | | WMSFE | | | | MVALPL | | | |
| | | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c |
| h= 1 | | 0.930*** | 0.916*** | 0.934*** | 0.935*** | 0.339*** | 0.925*** | 0.278*** | 0.285*** |
| h= 2 | | 0.934*** | 0.929** | 0.930*** | 0.926*** | 0.367*** | 1.021*** | 0.396*** | 0.401*** |
| h= 3 | | 0.948* | 0.944* | 0.946** | 0.940* | 0.353*** | 1.046*** | 0.349*** | 0.356*** |
| h= 6 | | 0.948 | 0.961 | 0.949 | 0.954 | 0.276* | 1.009*** | 0.251** | 0.296*** |
| h= 9 | | 0.945* | 0.957 | 0.946** | 0.960 | 0.241 | 1.017*** | 0.222* | 0.254*** |
| h=12 | | 0.982 | 0.996 | 0.983 | 0.994 | 0.110 | 0.886*** | 0.121 | 0.176*** |
| Intermediate VAR | | | | | | | | | |
| | | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c |
| h= 1 | | 0.919*** | 0.906*** | 0.940*** | 0.939*** | 0.358*** | 0.933*** | 0.270*** | 0.253*** |
| h= 2 | | 0.923*** | 0.919*** | 0.929*** | 0.924*** | 0.407*** | 1.011*** | 0.358*** | 0.360*** |
| h= 3 | | 0.934*** | 0.934** | 0.934*** | 0.939*** | 0.327*** | 1.023*** | 0.312*** | 0.264*** |
| h= 6 | | 0.933** | 0.937** | 0.936*** | 0.935** | 0.351*** | 1.054*** | 0.308*** | 0.276*** |
| h= 9 | | 0.939** | 0.939** | 0.940*** | 0.951** | 0.354*** | 1.043*** | 0.331*** | 0.271*** |
| h=12 | | 0.971** | 0.975 | 0.966*** | 0.965** | 0.194 | 0.956*** | 0.170 | 0.157* |
| Large VAR | | | | | | | | | |
| | | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c | BCVAR _{alt} | BCVAR | BCVAR _{c,alt} | BCVAR _c |
| h= 1 | | 0.907*** | 0.907*** | 0.918*** | 0.940*** | 0.410*** | 0.996*** | 0.365*** | 0.303*** |
| h= 2 | | 0.909*** | 0.909*** | 0.904*** | 0.908*** | 0.472*** | 1.139*** | 0.444*** | 0.406*** |
| h= 3 | | 0.914** | 0.916** | 0.908** | 0.922** | 0.423*** | 1.179*** | 0.437*** | 0.368*** |
| h= 6 | | 0.927 | 0.933 | 0.923* | 0.940 | 0.378*** | 1.131*** | 0.356*** | 0.269*** |
| h= 9 | | 0.928* | 0.938 | 0.923** | 0.943 | 0.337** | 1.076*** | 0.310** | 0.243*** |
| h=12 | | 0.951 | 0.969 | 0.951* | 0.968 | 0.231 | 1.009*** | 0.211 | 0.145 |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, {PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{BCVAR}_{\text{alt}}, \text{BCVAR}, \text{BCVAR}_{\text{c,alt}}, \text{BCVAR}_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.3. Out-of-sample forecast performance: Multivariate results, BVAR as the benchmark

| Fcst h. | | Medium VAR | | | | | | | | | |
|------------------|--|------------|--------------|-------|-----------------|--------------------|--------|--------|-------|---------------|--------------------|
| | | WMSFE | | | | | MVALPL | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.023 | 0.942* | 1.000 | 0.810*** | 0.826*** | -0.428 | -0.209 | 0.000 | -0.054 | -0.694 |
| h= 2 | | 0.943 | 0.943 | 1.000 | 0.833*** | 0.830*** | -0.235 | -0.249 | 0.000 | -0.047 | -0.666 |
| h= 3 | | 0.965 | 0.969 | 1.000 | 0.887* | 0.883* | -0.207 | -0.223 | 0.000 | -0.051 | -0.740 |
| h= 6 | | 1.010 | 0.975 | 1.000 | 0.944 | 0.938 | -0.162 | -0.193 | 0.000 | -0.021 | -0.734 |
| h= 9 | | 1.022 | 0.982 | 1.000 | 0.962 | 0.964 | -0.171 | -0.163 | 0.000 | -0.004 | -0.767 |
| h=12 | | 0.986 | 0.951 | 1.000 | 0.959 | 0.957 | -0.050 | -0.060 | 0.000 | -0.041 | -0.751 |
| Intermediate VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.052 | 0.951 | 1.000 | 0.821*** | 0.852*** | -0.278 | -0.168 | 0.000 | -0.055 | -0.735 |
| h= 2 | | 0.973 | 0.900** | 1.000 | 0.801*** | 0.805** | -0.048 | -0.051 | 0.000 | 0.116 | -0.536 |
| h= 3 | | 0.962 | 0.872** | 1.000 | 0.829** | 0.834* | -0.059 | -0.110 | 0.000 | 0.078 | -0.682 |
| h= 6 | | 1.040 | 1.004 | 1.000 | 0.960 | 0.958 | -0.250 | -0.359 | 0.000 | -0.133 | -0.911 |
| h= 9 | | 1.020 | 1.015 | 1.000 | 0.959 | 0.971 | -0.263 | -0.370 | 0.000 | -0.154 | -0.926 |
| h=12 | | 0.983 | 0.968 | 1.000 | 0.950 | 0.940 | -0.131 | -0.180 | 0.000 | -0.061 | -0.861 |
| Large VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.032 | 0.993 | 1.000 | 0.892*** | 0.925** | 0.045 | 0.030 | 0.000 | 0.091 | -0.602 |
| h= 2 | | 0.985 | 0.946 | 1.000 | 0.863*** | 0.862*** | 0.109 | 0.027 | 0.000 | 0.195 | -0.538 |
| h= 3 | | 0.985 | 0.927 | 1.000 | 0.876** | 0.882* | 0.074 | 0.025 | 0.000 | 0.205 | -0.606 |
| h= 6 | | 1.037 | 0.932 | 1.000 | 0.910* | 0.917 | 0.126 | 0.165 | 0.000 | 0.301 | -0.561 |
| h= 9 | | 1.039 | 0.956 | 1.000 | 0.930 | 0.934 | 0.092 | 0.075 | 0.000 | 0.197 | -0.636 |
| h=12 | | 1.003 | 0.938 | 1.000 | 0.924 | 0.923 | 0.225 | 0.201 | 0.000 | 0.300 | -0.564 |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark BVAR model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, {PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM, FAVAR, BVAR, BCVAR, BCVAR}_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark BVAR model, computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.4. Out-of-sample forecast performance: Multivariate results, alternative BIC

| Fcst h. | | Medium VAR | | | | | | | | | |
|------------------|--|------------|--------------|-------|-----------------|--------------------|----------|----------|-----------------|-----------------|--------------------|
| | | WMSFE | | | | | MVALPL | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.158 | 1.066 | 1.132 | 0.918*** | 0.930*** | 0.551*** | 0.770*** | 0.979*** | 0.906*** | 0.306*** |
| h= 2 | | 1.051 | 1.052 | 1.115 | 0.927** | 0.933** | 0.832*** | 0.818*** | 1.068*** | 1.031*** | 0.382*** |
| h= 3 | | 1.027 | 1.031 | 1.064 | 0.942* | 0.944* | 0.890*** | 0.874*** | 1.097*** | 1.048*** | 0.354*** |
| h= 6 | | 1.027 | 0.992 | 1.017 | 0.955 | 0.955 | 0.868*** | 0.837*** | 1.030*** | 1.002*** | 0.279*** |
| h= 9 | | 1.017 | 0.977 | 0.995 | 0.956 | 0.960 | 0.850*** | 0.858*** | 1.021*** | 0.976*** | 0.253*** |
| h=12 | | 1.025 | 0.988 | 1.039 | 1.001 | 0.995 | 0.877*** | 0.867*** | 0.927*** | 0.923*** | 0.192*** |
| Intermediate VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.160 | 1.048 | 1.103 | 0.907*** | 0.941*** | 0.710*** | 0.820*** | 0.988*** | 0.919*** | 0.245*** |
| h= 2 | | 1.117 | 1.033 | 1.148 | 0.921*** | 0.920*** | 0.847*** | 0.844*** | 0.895*** | 1.018*** | 0.346*** |
| h= 3 | | 1.083 | 0.981 | 1.126 | 0.936** | 0.931*** | 0.886*** | 0.835*** | 0.945*** | 1.023*** | 0.275*** |
| h= 6 | | 1.016 | 0.980** | 0.977 | 0.937** | 0.936** | 0.937*** | 0.828*** | 1.187*** | 1.064*** | 0.291*** |
| h= 9 | | 0.999 | 0.994 | 0.979 | 0.946** | 0.947** | 0.935*** | 0.828*** | 1.198*** | 1.056*** | 0.248*** |
| h=12 | | 1.009 | 0.993* | 1.026 | 0.974 | 0.967** | 0.886*** | 0.837*** | 1.017*** | 0.948*** | 0.176** |
| Large VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR | BCVAR _c | DFM | FAVAR | BVAR | BCVAR | BCVAR _c |
| h= 1 | | 1.049 | 1.009 | 1.017 | 0.901*** | 0.925*** | 0.950*** | 0.935*** | 0.905*** | 1.001*** | 0.323*** |
| h= 2 | | 1.037 | 0.996 | 1.053 | 0.905*** | 0.905*** | 1.053*** | 0.971*** | 0.944*** | 1.124*** | 0.390*** |
| h= 3 | | 1.030 | 0.970 | 1.045 | 0.914** | 0.905** | 1.049*** | 0.999*** | 0.974*** | 1.139*** | 0.369*** |
| h= 6 | | 1.063 | 0.955 | 1.026 | 0.934 | 0.927 | 0.957*** | 0.995*** | 0.830*** | 1.153*** | 0.306*** |
| h= 9 | | 1.049 | 0.965 | 1.009 | 0.938 | 0.926* | 0.972*** | 0.954*** | 0.879*** | 1.089*** | 0.263*** |
| h=12 | | 1.052 | 0.984 | 1.049 | 0.974 | 0.948 | 0.934*** | 0.910*** | 0.709** | 0.992*** | 0.144 |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, $\{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM}, \text{FAVAR}, \text{BVAR}, \text{BCVAR}, \text{BCVAR}_c\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.5. Out-of-sample forecast performance: Multivariate results, 7 key variables ordered last

| Fest h. | | Medium VAR | | | | | | | | | |
|------------------|--|------------|--------------|-------|--------------------|------------------------|-----------------|-----------------|-----------------|--------------------|------------------------|
| | | WMSFE | | | | | MVALPL | | | | |
| | | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} |
| h= 1 | | 1.158 | 1.066 | 1.132 | 0.935*** | 0.935*** | 0.551*** | 0.770*** | 0.979*** | 0.285*** | 0.301*** |
| h= 2 | | 1.051 | 1.052 | 1.115 | 0.926*** | 0.926*** | 0.832*** | 0.818*** | 1.068*** | 0.401*** | 0.403*** |
| h= 3 | | 1.027 | 1.031 | 1.064 | 0.940* | 0.939** | 0.890*** | 0.874*** | 1.097*** | 0.356*** | 0.357*** |
| h= 6 | | 1.027 | 0.992 | 1.017 | 0.954 | 0.956 | 0.868*** | 0.837*** | 1.030*** | 0.296*** | 0.279*** |
| h= 9 | | 1.017 | 0.977 | 0.995 | 0.960 | 0.953 | 0.850*** | 0.858*** | 1.021*** | 0.254*** | 0.237*** |
| h=12 | | 1.025 | 0.988 | 1.039 | 0.994 | 0.995 | 0.877*** | 0.867*** | 0.927*** | 0.176*** | 0.166** |
| Intermediate VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} |
| h= 1 | | 1.160 | 1.048 | 1.103 | 0.939*** | 0.936*** | 0.710*** | 0.820*** | 0.988*** | 0.253*** | 0.293*** |
| h= 2 | | 1.117 | 1.033 | 1.148 | 0.924*** | 0.945*** | 0.847*** | 0.844*** | 0.895*** | 0.360*** | 0.313*** |
| h= 3 | | 1.083 | 0.981 | 1.126 | 0.939*** | 0.957** | 0.886*** | 0.835*** | 0.945*** | 0.264*** | 0.267*** |
| h= 6 | | 1.016 | 0.980** | 0.977 | 0.935** | 0.952** | 0.937*** | 0.828*** | 1.187*** | 0.276*** | 0.242*** |
| h= 9 | | 0.999 | 0.994 | 0.979 | 0.951** | 0.965* | 0.935*** | 0.828*** | 1.198*** | 0.271*** | 0.264*** |
| h=12 | | 1.009 | 0.993* | 1.026 | 0.965** | 0.977 | 0.886*** | 0.837*** | 1.017*** | 0.157* | 0.142 |
| Large VAR | | | | | | | | | | | |
| | | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} | DFM | FAVAR | BVAR | BCVAR _c | BCVAR _{c,v.2} |
| h= 1 | | 1.049 | 1.009 | 1.017 | 0.940*** | 0.945*** | 0.950*** | 0.935*** | 0.905*** | 0.303*** | 0.304*** |
| h= 2 | | 1.037 | 0.996 | 1.053 | 0.908*** | 0.940* | 1.053*** | 0.971*** | 0.944*** | 0.406*** | 0.371*** |
| h= 3 | | 1.030 | 0.970 | 1.045 | 0.922** | 0.947 | 1.049*** | 0.999*** | 0.974*** | 0.368*** | 0.317*** |
| h= 6 | | 1.063 | 0.955 | 1.026 | 0.940 | 0.953 | 0.957*** | 0.995*** | 0.830*** | 0.269*** | 0.267*** |
| h= 9 | | 1.049 | 0.965 | 1.009 | 0.943 | 0.961 | 0.972*** | 0.954*** | 0.879*** | 0.243*** | 0.217** |
| h=12 | | 1.052 | 0.984 | 1.049 | 0.968 | 0.973 | 0.934*** | 0.910*** | 0.709** | 0.145 | 0.143* |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, {PAYEMS, CPIAUCSL, FEDFUNDS, INDPRO, UNRATE, PPIFGS, GS10}. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{DFM, FAVAR, BVAR, BCVAR}_c, \text{BCVAR}_{c,v.2}\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.6. Out-of-sample forecast performance: Compressed TVP-SV VAR, BVAR as the benchmark

| Variable | Medium VAR | | | | | | | | | | | | |
|--------------|------------------|-----------------|-----------------|----------------|-----------------|-----------------|-----------------|-----------------|-----------------|---------------|--------------|---------------|---------------|
| | MSFE | | | | | | ALPL | | | | | | |
| | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 | | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 |
| PAYEMS | 0.810** | 1.020 | 1.082 | 0.948 | 0.933 | 0.937 | 0.120*** | 0.025 | -0.012 | -0.167 | -0.514 | -0.573 | |
| CPIAUCSL | 0.973 | 0.873** | 0.848*** | 0.823** | 0.811*** | 0.790*** | 0.959 | 1.880 | 1.444 | 1.051 | 1.026 | 1.197 | |
| FEDFUNDS | 0.319*** | 0.364*** | 0.497** | 0.833* | 0.975 | 0.985 | 0.629*** | 0.479* | 0.308 | 0.263 | 0.184 | 0.256 | |
| INDPRO | 1.109 | 1.121 | 1.009 | 0.956 | 0.957 | 0.985 | 0.068 | -0.175 | -0.127 | -0.281 | -0.262 | -0.237 | |
| UNRATE | 1.081 | 1.053 | 1.031 | 0.992 | 0.999 | 1.003 | -0.043 | -0.026 | -0.014 | 0.001 | -0.003 | -0.042 | |
| PPIFGS | 0.987 | 0.914* | 0.908*** | 0.879* | 0.889* | 0.886* | 0.718** | 1.074 | 0.884* | 1.091 | 0.820 | 0.744 | |
| GS10 | 0.932 | 0.941 | 0.912 | 0.924 | 0.971 | 0.979 | 0.011 | -0.007 | -0.063 | -0.060 | -0.046 | 0.020 | |
| Multivariate | 0.800*** | 0.793*** | 0.838** | 0.901* | 0.928** | 0.931** | 0.674*** | 0.634*** | 0.476*** | 0.194 | 0.028 | -0.077 | |
| | Intermediate VAR | | | | | | Large VAR | | | | | | |
| | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 | |
| PAYEMS | 0.883 | 1.106 | 1.211 | 1.072 | 0.970 | 0.909 | 0.072** | -0.019 | -0.081 | -0.364 | -0.674 | -0.992 | |
| CPIAUCSL | 0.939 | 0.791* | 0.752* | 0.853* | 0.875** | 0.900* | 1.043** | 2.586 | 2.179 | 1.070 | 1.217 | 0.902 | |
| FEDFUNDS | 0.357*** | 0.334** | 0.388** | 0.755** | 0.821** | 0.806* | 0.692*** | 0.642** | 0.502 | 0.221 | -0.072 | 0.174 | |
| INDPRO | 1.162 | 1.206 | 1.098 | 0.981 | 0.981 | 0.942 | -0.041 | -0.266 | -0.133 | -0.033 | -0.113 | -0.211 | |
| UNRATE | 1.118 | 1.087 | 1.070 | 1.044 | 0.994 | 0.981 | -0.066 | -0.046 | -0.041 | -0.040 | -0.019 | -0.030 | |
| PPIFGS | 0.935 | 0.845* | 0.842** | 0.907 | 0.913* | 0.937* | 0.963*** | 1.505* | 1.458* | 1.152 | 1.026 | 0.908 | |
| GS10 | 0.910 | 0.902 | 0.856*** | 0.932** | 0.971 | 0.979 | 0.041 | -0.007 | -0.021 | -0.011 | -0.047 | -0.030 | |
| Multivariate | 0.825*** | 0.765*** | 0.779** | 0.918** | 0.927** | 0.917** | 0.645*** | 0.739*** | 0.566** | 0.027 | -0.232 | -0.344 | |
| | Large VAR | | | | | | | | | | | | |
| | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 | <i>h</i> = 1 | <i>h</i> = 2 | <i>h</i> = 3 | <i>h</i> = 6 | <i>h</i> = 9 | <i>h</i> = 12 | |
| PAYEMS | 0.915 | 1.177 | 1.157 | 1.059 | 1.026 | 1.010 | 0.036 | -0.066 | -0.073 | -0.213 | -0.574 | -0.607 | |
| CPIAUCSL | 1.052 | 0.907* | 0.863** | 0.813** | 0.783** | 0.768* | 0.603* | 2.482 | 2.656 | 2.539 | 1.534 | 1.735 | |
| FEDFUNDS | 0.429*** | 0.418*** | 0.505** | 0.771* | 0.928 | 0.934 | 0.424** | 0.330 | 0.261 | 0.258 | -0.175 | 0.058 | |
| INDPRO | 1.152 | 1.159 | 1.072 | 0.974 | 0.938 | 0.950 | 0.427 | 0.094 | -0.249 | -0.169 | -0.022 | -0.036 | |
| UNRATE | 1.050 | 1.106 | 1.053 | 1.071 | 1.045 | 1.044 | -0.003 | -0.061 | -0.028 | -0.034 | -0.011 | -0.024 | |
| PPIFGS | 1.048 | 0.927 | 0.923** | 0.875** | 0.866** | 0.856* | 1.282** | 2.176* | 1.686* | 1.940 | 1.612 | 0.937 | |
| GS10 | 0.926 | 0.899 | 0.850** | 0.868** | 0.943 | 0.944 | 0.001 | 0.046 | 0.044 | 0.033 | -0.010 | -0.067 | |
| Multivariate | 0.887*** | 0.838*** | 0.847** | 0.899* | 0.923** | 0.922* | 0.761*** | 0.722** | 0.618** | 0.386* | 0.123 | 0.004 | |

The left half of this table reports the ratio between the univariate or multivariate weighted mean squared forecast error of the BCVAR_{*tvp-sv*} model and the univariate or multivariate weighted mean squared forecast error of the benchmark BVAR model. The right half of the table shows the univariate or multivariate average log predictive likelihood differentials between the BCVAR_{*tvp-sv*} model and the benchmark BVAR model. *h* denotes the forecast horizons, with $h \in \{1, 2, 3, 6, 9, 12\}$. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate all instances where the BCVAR_{*tvp-sv*} model outperforms all alternative models (DFM, FAVAR, BVAR, BCVAR, BCVAR_c), for any given VAR size/variable/forecast horizon combination. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.7. Out-of-sample forecast performance: Compressed TVP-SV VAR, alternative BIC

| Variable | Medium VAR | | | | | | | | | | | | |
|------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|---------------|----------|
| | MSFE | | | | | | ALPL | | | | | | |
| | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ | | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ |
| PAYEMS | 0.742*** | 0.586*** | 0.572*** | 0.646** | 0.773* | 0.890 | 0.323*** | 0.384*** | 0.324*** | -0.032 | -0.557 | -0.759 | |
| CPIAUCSL | 0.919*** | 0.874*** | 0.889*** | 0.860** | 0.836*** | 0.850*** | 0.278* | 0.247** | 0.416 | 0.159* | 0.270 | 0.190* | |
| FEDFUNDS | 0.889 | 0.872 | 0.947 | 1.024 | 1.001 | 1.106 | 0.601** | 0.635*** | 0.539* | 0.320 | 0.321 | 0.193 | |
| INDPRO | 0.911** | 0.933 | 0.974 | 0.981 | 0.992 | 0.991 | -0.021 | -0.100 | -0.170 | -0.614 | -0.477 | -0.186 | |
| UNRATE | 0.839*** | 0.847** | 0.885* | 0.937 | 0.982 | 1.020 | 0.113*** | 0.105*** | 0.086*** | 0.051** | 0.038 | -0.002 | |
| PPIFGS | 0.985 | 0.995 | 1.014 | 0.998 | 0.996 | 1.012 | 0.277** | 0.348 | 0.386 | 0.334 | 0.369 | 0.324 | |
| GS10 | 1.022 | 1.033 | 1.047 | 1.036 | 1.012 | 1.038 | 0.007 | -0.009 | -0.089 | -0.049 | -0.016 | -0.059 | |
| Multivariate | 0.913*** | 0.888*** | 0.906** | 0.919* | 0.934 | 0.979 | 1.581*** | 1.703*** | 1.527*** | 1.144*** | 1.008*** | 0.819** | |
| Intermediate VAR | | | | | | | | | | | | | |
| | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ | | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ |
| PAYEMS | 0.741*** | 0.586*** | 0.594*** | 0.655*** | 0.736** | 0.848 | 0.313*** | 0.386*** | 0.251** | -0.149 | -0.614 | -0.957 | |
| CPIAUCSL | 0.943 | 0.869*** | 0.873*** | 0.844*** | 0.782*** | 0.819*** | 0.224 | 0.427 | 0.326 | 0.316 | 0.193*** | 0.247 | |
| FEDFUNDS | 0.878** | 0.853** | 0.852* | 0.957 | 0.986 | 1.065 | 0.689*** | 0.655*** | 0.633*** | 0.562*** | 0.214 | 0.397 | |
| INDPRO | 0.920*** | 0.928** | 0.933* | 0.969 | 0.959 | 0.984 | 0.069 | -0.124 | -0.174 | -0.389 | -0.258 | -0.301 | |
| UNRATE | 0.870*** | 0.895** | 0.915** | 0.925** | 0.971 | 0.986 | 0.105*** | 0.092*** | 0.080*** | 0.051*** | 0.017 | -0.009 | |
| PPIFGS | 0.974 | 1.005 | 0.986 | 0.970 | 0.962 | 0.979 | 0.258 | 0.363 | 0.338 | 0.361 | 0.355 | 0.311 | |
| GS10 | 1.017 | 1.021 | 1.045 | 1.017 | 1.014 | 1.012 | -0.013 | -0.038 | -0.048 | -0.049 | 0.006 | -0.034 | |
| Multivariate | 0.919*** | 0.891*** | 0.887*** | 0.899*** | 0.907*** | 0.949** | 1.569*** | 1.572*** | 1.457*** | 1.160*** | 0.972*** | 0.711* | |
| Large VAR | | | | | | | | | | | | | |
| | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ | | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ |
| PAYEMS | 0.718*** | 0.569*** | 0.557*** | 0.658* | 0.768 | 0.879 | 0.329*** | 0.396*** | 0.340*** | 0.069 | -0.279 | -0.547 | |
| CPIAUCSL | 0.923* | 0.858*** | 0.862*** | 0.863* | 0.816*** | 0.811*** | 0.260* | 0.431* | 0.442 | 0.347 | 0.409 | 0.248 | |
| FEDFUNDS | 0.904 | 0.891 | 0.905 | 1.007 | 1.016 | 1.074 | 0.648** | 0.642** | 0.447 | 0.469 | 0.322 | 0.290 | |
| INDPRO | 0.908*** | 0.929* | 0.940 | 1.003 | 1.016 | 1.016 | -0.038 | -0.204 | -0.043 | -0.252 | -0.250 | -0.400 | |
| UNRATE | 0.850*** | 0.863* | 0.901 | 0.961 | 0.978 | 1.027 | 0.115*** | 0.096*** | 0.076*** | 0.053** | 0.035 | 0.007 | |
| PPIFGS | 0.980 | 1.002 | 0.997 | 1.001 | 1.004 | 1.018 | 0.250* | 0.344 | 0.370 | 0.329 | 0.403 | 0.228 | |
| GS10 | 1.023 | 1.032 | 1.036 | 1.044 | 1.018 | 1.018 | -0.002 | 0.003 | -0.005 | -0.025 | 0.011 | -0.022 | |
| Multivariate | 0.914*** | 0.889*** | 0.887** | 0.927 | 0.937 | 0.969 | 1.629*** | 1.690*** | 1.562*** | 1.224*** | 1.023*** | 0.700* | |

The left half of this table reports the ratio between the univariate or multivariate weighted mean squared forecast error of the BCVAR_{tvp-sv} model and the univariate or multivariate weighted mean squared forecast error of the benchmark AR(1) model. The right half of the table shows the univariate or multivariate average log predictive likelihood differentials between the BCVAR_{tvp-sv} model and the benchmark AR(1) model. h denotes the forecast horizons, with $h \in \{1, 2, 3, 6, 9, 12\}$. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate all instances where the BCVAR_{tvp-sv} model outperforms all alternative models (DFM, FAVAR, BVAR, BCVAR, BCVAR_c), for any given VAR size/variable/forecast horizon combination. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.8. Out-of-sample forecast performance: Compressed TVP-SV VAR, 7 key variables ordered last

| Variable | Medium VAR | | | | | | | | | | | |
|------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|----------------|--------------|
| | MSFE | | | | | | ALPL | | | | | |
| | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ | $h = 1$ | $h = 2$ | $h = 3$ | $h = 6$ | $h = 9$ | $h = 12$ |
| PAYEMS | 0.727*** | 0.584*** | 0.564*** | 0.658** | 0.775* | 0.878 | 0.322*** | 0.380*** | 0.337*** | 0.043 | -0.351 | -0.546 |
| CPIAUCSL | 0.906*** | 0.875*** | 0.871*** | 0.861*** | 0.830*** | 0.853*** | 0.259* | 0.327** | 0.442 | 0.138** | 0.317* | 0.117 |
| FEDFUNDS | 0.906 | 0.872 | 0.922 | 0.997 | 0.969 | 1.050 | 0.699*** | 0.480 | 0.560** | 0.447 | 0.105 | 0.263 |
| INDPRO | 0.891*** | 0.929* | 0.934 | 0.967 | 0.991 | 0.988 | -0.045 | -0.177 | -0.264 | -0.480 | -0.443 | -0.196 |
| UNRATE | 0.879*** | 0.879* | 0.895* | 0.961 | 0.968 | 1.003 | 0.105*** | 0.095*** | 0.087*** | 0.054*** | 0.047** | 0.010 |
| PPIFGS | 0.979 | 0.994 | 0.990 | 0.995 | 0.999 | 1.006 | 0.282** | 0.351 | 0.388 | 0.323 | 0.390 | 0.323 |
| GS10 | 1.012 | 1.023 | 1.032 | 1.013 | 0.993 | 1.022 | -0.039 | -0.009 | -0.021 | -0.060 | 0.019 | -0.010 |
| Multivariate | 0.912*** | 0.891*** | 0.889*** | 0.915* | 0.925* | 0.965 | 1.634*** | 1.659*** | 1.566*** | 1.173*** | 0.990*** | 0.799** |
| Intermediate VAR | | | | | | | | | | | | |
| PAYEMS | 0.771*** | 0.600*** | 0.591*** | 0.644** | 0.748** | 0.868 | 0.301*** | 0.360*** | 0.292*** | 0.128 | -0.340 | -0.969 |
| CPIAUCSL | 0.949 | 0.882*** | 0.874*** | 0.849*** | 0.805*** | 0.813*** | 0.251* | 0.456 | 0.342 | 0.284 | 0.273 | 0.227 |
| FEDFUNDS | 0.893** | 0.834** | 0.862* | 0.937 | 1.024 | 1.042 | 0.840*** | 0.700*** | 0.528* | 0.439 | 0.072 | 0.203 |
| INDPRO | 0.895*** | 0.962 | 0.975 | 0.988 | 0.989 | 1.004 | 0.028 | -0.005 | -0.110 | -0.359 | -0.258 | -0.278 |
| UNRATE | 0.873*** | 0.899** | 0.946 | 0.935** | 0.970 | 1.004 | 0.098*** | 0.089*** | 0.056** | 0.038* | 0.014 | -0.003 |
| PPIFGS | 0.961 | 0.978 | 0.970 | 0.954 | 0.947** | 0.977 | 0.274 | 0.404 | 0.390 | 0.342 | 0.375 | 0.289 |
| GS10 | 1.012 | 1.045 | 1.041 | 1.029 | 1.011 | 1.005 | 0.041** | 0.010 | -0.052 | 0.008 | -0.019 | -0.010 |
| Multivariate | 0.918*** | 0.897*** | 0.895*** | 0.899*** | 0.919** | 0.952* | 1.602*** | 1.551*** | 1.458*** | 1.167*** | 0.910*** | 0.674* |
| Large VAR | | | | | | | | | | | | |
| PAYEMS | 0.724*** | 0.601*** | 0.566*** | 0.683* | 0.785 | 0.898 | 0.310*** | 0.384*** | 0.338*** | 0.091 | -0.259 | -0.504 |
| CPIAUCSL | 0.931* | 0.864*** | 0.903** | 0.859** | 0.795*** | 0.813*** | 0.229 | 0.348* | 0.307 | 0.300 | 0.347 | 0.154 |
| FEDFUNDS | 0.959 | 0.929 | 0.938 | 1.020 | 1.028 | 1.085 | 0.628** | 0.574** | 0.523 | 0.513* | 0.367 | -0.220 |
| INDPRO | 0.904** | 0.969 | 0.955 | 1.041 | 1.015 | 1.038 | 0.162* | -0.133 | -0.167 | -0.158 | -0.349 | -0.338 |
| UNRATE | 0.876** | 0.879* | 0.900 | 0.947 | 1.020 | 1.013 | 0.103*** | 0.087*** | 0.081*** | 0.048** | -0.006 | 0.018 |
| PPIFGS | 0.973 | 1.021 | 0.990 | 1.027 | 0.977 | 1.003 | 0.219 | 0.331 | 0.360 | 0.265 | 0.386 | 0.251 |
| GS10 | 1.028 | 1.039 | 1.062 | 1.045 | 1.021 | 1.043 | 0.001 | 0.016 | -0.057 | -0.035 | -0.090 | -0.042 |
| Multivariate | 0.925*** | 0.910*** | 0.904** | 0.939 | 0.940 | 0.977 | 1.554*** | 1.570*** | 1.485*** | 1.138*** | 0.910*** | 0.714** |

The left half of this table reports the ratio between the univariate or multivariate weighted mean squared forecast error of the BCVAR_{tvp-sv} model and the univariate or multivariate weighted mean squared forecast error of the benchmark AR(1) model. The right half of the table shows the univariate or multivariate average log predictive likelihood differentials between the BCVAR_{tvp-sv} model and the benchmark AR(1) model. h denotes the forecast horizons, with $h \in \{1, 2, 3, 6, 9, 12\}$. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate all instances where the BCVAR_{tvp-sv} model outperforms all alternative models (DFM, FAVAR, BVAR, BCVAR, BCVAR_c), for any given VAR size/variable/forecast horizon combination. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Table C.9. Out-of-sample forecast performance: Multivariate results, alternative SV models

| Fcst h . | | Small VAR | | | MVALPL | | |
|------------------|--|---------------------|---------------------|-------------------------|---------------------|---------------------|-------------------------|
| | | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} |
| h= 1 | | 0.917*** | 0.942*** | 0.918*** | 2.047*** | 1.696*** | 1.719*** |
| h= 2 | | 0.930*** | 0.944*** | 0.895*** | 1.907*** | 1.654*** | 1.745*** |
| h= 3 | | 0.936*** | 0.951** | 0.901*** | 1.845*** | 1.563*** | 1.645*** |
| h= 6 | | 0.946*** | 0.971 | 0.912*** | 1.608*** | 1.228*** | 1.386*** |
| h= 9 | | 0.968*** | 0.981 | 0.936*** | 1.385*** | 0.978*** | 1.143*** |
| h=12 | | 0.992 | 0.999 | 0.960* | 0.931* | 0.811* | 0.930*** |
| Medium VAR | | | | | | | |
| | | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} |
| h= 1 | | 1.070 | 0.935*** | 0.905*** | 1.599*** | 1.522*** | 1.653*** |
| h= 2 | | 1.089 | 0.922*** | 0.884*** | 1.521*** | 1.558*** | 1.701*** |
| h= 3 | | 1.123 | 0.931** | 0.892*** | 1.236*** | 1.399*** | 1.573*** |
| h= 6 | | 1.125 | 0.937* | 0.916* | 1.041*** | 1.129*** | 1.224*** |
| h= 9 | | 1.031 | 0.947* | 0.924* | 1.078*** | 0.938*** | 1.049*** |
| h=12 | | 1.007 | 0.981 | 0.967 | 1.039*** | 0.760** | 0.851*** |
| Intermediate VAR | | | | | | | |
| | | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} |
| h= 1 | | | 0.942*** | 0.910*** | | 1.520*** | 1.633*** |
| h= 2 | | | 0.924*** | 0.878*** | | 1.512*** | 1.635*** |
| h= 3 | | | 0.927*** | 0.877*** | | 1.324*** | 1.511*** |
| h= 6 | | | 0.919** | 0.896*** | | 1.119*** | 1.215*** |
| h= 9 | | | 0.932** | 0.908*** | | 0.895*** | 0.966*** |
| h=12 | | | 0.959** | 0.941** | | 0.718* | 0.674 |
| Large VAR | | | | | | | |
| | | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} | BVAR _{ccm} | BCVAR _{sv} | BCVAR _{tvp-sv} |
| h= 1 | | | 0.942*** | 0.902*** | | 1.488*** | 1.667*** |
| h= 2 | | | 0.924** | 0.883*** | | 1.543*** | 1.666*** |
| h= 3 | | | 0.919** | 0.885** | | 1.394*** | 1.593*** |
| h= 6 | | | 0.939 | 0.922 | | 1.118*** | 1.216*** |
| h= 9 | | | 0.938 | 0.932 | | 0.894*** | 1.002*** |
| h=12 | | | 0.950 | 0.967 | | 0.722* | 0.713* |

The left half of this table reports the ratio between the multivariate weighted mean squared forecast error (WMSFE) of model i and the WMSFE of the benchmark AR(1) model, computed as

$$WMSFE_{ih} = \frac{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{i,\tau+h}}{\sum_{\tau=\underline{t}}^{\bar{t}-h} we_{bcmk,\tau+h}},$$

where $we_{i,\tau+h} = (e'_{i,\tau+h} \times W \times e_{i,\tau+h})$ and $we_{bcmk,\tau+h} = (e'_{bcmk,\tau+h} \times W \times e_{bcmk,\tau+h})$ denote the weighted forecast errors of model i and the benchmark model at time $\tau+h$, $e_{i,\tau+h}$ and $e_{bcmk,\tau+h}$ are the $(N \times 1)$ vector of forecast errors, and W is an $(N \times N)$ matrix of weights. We set $N = 7$, to focus on the following key seven series, $\{\text{PAYEMS}, \text{CPIAUCSL}, \text{FEDFUNDS}, \text{INDPRO}, \text{UNRATE}, \text{PPIFGS}, \text{GS10}\}$. In addition, we set the matrix W to be a diagonal matrix featuring on the diagonal the inverse of the variances of the series to be forecast. \underline{t} and \bar{t} denote the start and end of the out-of-sample period, $i \in \{\text{BVAR}_{ccm}, \text{BCVAR}_{sv}, \text{BCVAR}_{tvp-sv}\}$, and $h \in \{1, 2, 3, 6, 9, 12\}$. The right half of the table shows the multivariate average log predictive likelihood differentials between model i and the benchmark AR(1), computed as

$$MVALPL_{ih} = \frac{1}{\bar{t} - \underline{t} - h + 1} \sum_{\tau=\underline{t}}^{\bar{t}-h} (MVLPL_{i,\tau+h} - MVLPL_{bcmk,\tau+h}),$$

where $MVLPL_{i,\tau+h}$ and $MVLPL_{bcmk,\tau+h}$ denote the multivariate log predictive likelihoods of model i and the benchmark model at time $\tau+h$, and are computed under the assumption of joint normality. All forecasts are generated out-of-sample using recursive estimates of the models, with the out of sample period starting in 1987:07 and ending in 2014:12. Bold numbers indicate the lowest WMSFE and highest MVALPL across all models for any given VAR size - forecast horizon pair. * significance at the 10% level; ** significance at the 5% level; *** significance at the 1% level.

Appendix D Data and transformations

Table D.1. Output and Income

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|---------------|-------------------------------------|------------------------|
| 1 | 5 | X | X | RPI | Real Personal Income | PI |
| 2 | 5 | | X | W875RX1 | RPI ex. Transfers | PI less transfers |
| 6 | 5 | X | X | INDPRO | IP Index | IP: total |
| 7 | 5 | | | IPFPNSS | IP: Final Products and Supplies | IP: products |
| 8 | 5 | | | IPFINAL | IP: Final Products | IP: final prod |
| 9 | 5 | | | IPCONGD | IP: Consumer Goods | IP: cons gds |
| 10 | 5 | | | IPDCONGD | IP: Durable Consumer Goods | IP: cons dble |
| 11 | 5 | | | IPNCONGD | IP: Nondurable Consumer Goods | IP: cons nondble |
| 12 | 5 | | | IPBUSEQ | IP: Business Equipment | IP: bus eqpt |
| 13 | 5 | | | IPMAT | IP: Materials | IP: matls |
| 14 | 5 | | | IPDMAT | IP: Durable Materials | IP: dble matls |
| 15 | 5 | | | IPNMAT | IP: Nondurable Materials | IP: nondble matls |
| 16 | 5 | | | IPMANSICS | IP: Manufacturing | IP: mfg |
| 17 | 5 | | | IPB51222S | IP: Residential Utilities | IP: res util |
| 18 | 5 | | | IPFUELS | IP: Fuels | IP: fuels |
| 19 | 1 | | X | NAPMPI | ISM Manufacturing: Production | NAPM prodn |
| 20 | 1 | | | CAPUTLB00004S | Capacity Utilization: Manufacturing | Cap util |

Table D.2. Labor Market

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|---------------|-------------------------------------|------------------------|
| 21 | 1 | X | X | HWI | Help-Wanted Index for US | Help wanted indx |
| 22 | 1 | | X | HWIURATIO | Help Wanted to Unemployed ratio | Help wanted/unemp |
| 23 | 5 | | X | CLF16OV | Civilian Labor Force | Emp CPS total |
| 24 | 5 | | | CE16OV | Civilian Employment | Emp CPS nonag |
| 25 | 2 | X | X | UNRATE | Civilian Unemployment Rate | U: all |
| 26 | 1 | | | UEMPMEAN | Average Duration of Unemployment | U: mean duration |
| 27 | 5 | | | UEMPLT5 | Civilians Unemployed ≤ 5 Weeks | U ≤ 5 wks |
| 28 | 5 | | | UEMP5TO14 | Civilians Unemployed 5-14 Weeks | U 5-14 wks |
| 29 | 5 | | | UEMP15OV | Civilians Unemployed > 15 Weeks | U > 15 wks |
| 30 | 5 | | | UEMP15T26 | Civilians Unemployed 15-26 Weeks | U 15-26 wks |
| 31 | 5 | | | UEMP27OV | Civilians Unemployed > 27 Weeks | U > 27 wks |
| 32 | 5 | | | CLAIMSx | Initial Claims | UI claims |
| 33 | 5 | X | X | PAYEMS | All Employees: Total nonfarm | Emp: total |
| 34 | 5 | | | USGOOD | All Employees: Goods-Producing | Emp: gds prod |
| 35 | 5 | | | CES1021000001 | All Employees: Mining and Logging | Emp: mining |
| 36 | 5 | | | USCONS | All Employees: Construction | Emp: const |
| 37 | 5 | | | MANEMP | All Employees: Manufacturing | Emp: mfg |
| 38 | 5 | | | DMANEMP | All Employees: Durable goods | Emp: dble gds |
| 39 | 5 | | | NDMANEMP | All Employees: Nondurable goods | Emp: nondbles |
| 40 | 5 | | | SRVPRD | All Employees: Service Industries | Emp: services |
| 41 | 5 | | | USTPU | All Employees: TT&U | Emp: TTU |
| 42 | 5 | | | USWTRADE | All Employees: Wholesale Trade | Emp: wholesale |
| 43 | 5 | | | USTRADE | All Employees: Retail Trade | Emp: retail |
| 44 | 5 | | | USFIRE | All Employees: Financial Activities | Emp: FIRE |
| 45 | 5 | | | USGOVT | All Employees: Government | Emp: Govt |
| 46 | 1 | | X | CES0600000007 | Hours: Goods-Producing | Avg hrs |
| 47 | 1 | | | AWOTMAN | Overtime Hours: Manufacturing | Overtime: mfg |
| 48 | 1 | | | AWHMAN | Hours: Manufacturing | Avg hrs: mfg |
| 49 | 1 | | | NAPMEI | ISM Manufacturing: Employment | NAPM empl |
| 128 | 5 | | | CES0600000008 | Ave. Hourly Earnings: Goods | AHE: goods |
| 129 | 5 | | | CES2000000008 | Ave. Hourly Earnings: Construction | AHE: const |
| 130 | 5 | | | CES3000000008 | Ave. Hourly Earnings: Manufacturing | AHE: mfg |

Table D.3. **Housing**

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|-------------|--------------------|------------------------|
| 50 | 4 | X | | HOUST | Starts: Total | Starts: nonfarm |
| 51 | 4 | | | HOUSTNE | Starts: Northeast | Starts: NE |
| 52 | 4 | | | HOUSTMW | Starts: Midwest | Starts: MW |
| 53 | 4 | | | HOUSTS | Starts: South | Starts: South |
| 54 | 4 | | | HOUSTW | Starts: West | Starts: West |
| 55 | 4 | X | | PERMIT | Permits | BP: total |
| 56 | 4 | | | PERMITNE | Permits: Northeast | BP: NE |
| 57 | 4 | | | PERMITMW | Permits: Midwest | BP: MW |
| 58 | 4 | | | PERMITS | Permits: South | BP: South |
| 59 | 4 | | | PERMITW | Permits: West | BP: West |

Table D.4. **Consumption, Orders and Inventories**

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|-----------------|--------------------------------|------------------------|
| 3 | 5 | | X | DPCERA3M086SBEA | Real PCE | Real Consumption |
| 4 | 5 | | X | CMRMTSPLx | Real M&T Sales | M&T sales |
| 5 | 5 | | X | RETAILx | Retail and Food Services Sales | Retail sales |
| 60 | 1 | X | | NAPM | ISM: PMI Composite Index | PMI |
| 61 | 1 | | X | NAPMNOI | ISM: New Orders Index | NAPM new ordrs |
| 62 | 1 | | X | NAPMSDI | ISM: Supplier Deliveries Index | NAPM vendor del |
| 63 | 1 | | X | NAPMII | ISM: Inventories Index | NAPM Invent |
| 65 | 5 | | | AMDMNOx | Orders: Durable Goods | Orders: dble gds |
| 67 | 5 | | | AMDMUOx | Unfilled Orders: Durable Goods | Unf orders: dble |
| 68 | 5 | | | BUSINVx | Total Business Inventories | M&T invent |
| 69 | 1 | | | ISRATIOx | Inventories to Sales Ratio | M&T invent/sales |

Table D.5. **Money and Credit**

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|-------------|----------------------------------|------------------------|
| 70 | 5 | X | X | M1SL | M1 Money Stock | M1 |
| 71 | 5 | | X | M2SL | M2 Money Stock | M2 |
| 73 | 5 | | X | M2REAL | Real M2 Money Stock | M2 (real) |
| 74 | 5 | | X | AMBSL | St. Louis Adjusted Monetary Base | MB |
| 75 | 5 | | X | TOTRESNS | Total Reserves | Reserves tot |
| 77 | 5 | X | X | BUSLOANS | Commercial and Industrial Loans | C&I loan plus |
| 78 | 5 | | | REALLN | Real Estate Loans | DC&I loans |
| 79 | 5 | | X | NONREVSL | Total Nonrevolving Credit | Cons credit |
| 80 | 1 | | X | CONSPI | Credit to PI ratio | Inst cred/PI |
| 132 | 5 | | | MZMSL | MZM Money Stock | N.A. |
| 133 | 5 | | | DTCOLNVHFNM | Consumer Motor Vehicle Loans | N.A. |
| 134 | 5 | | | DTCTHFNM | Total Consumer Loans and Leases | N.A. |
| 135 | 5 | X | | INVEST | Securities in Bank Credit | N.A. |

Table D.6. Interest rates and Exchange rates

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|-------------|------------------------------|------------------------|
| 85 | 2 | X | X | FEDFUNDS | Effective Federal Funds Rate | Fed Funds |
| 86 | 2 | | X | CP3M | 3-Month AA Comm. Paper Rate | Comm paper |
| 87 | 2 | | X | TB3MS | 3-Month T-bill | 3 mo T-bill |
| 88 | 2 | | X | TB6MS | 6-Month T-bill | 6 mo T-bill |
| 89 | 2 | | X | GS1 | 1-Year T-bond | 1 yr T-bond |
| 90 | 2 | | X | GS5 | 5-Year T-bond | 5 yr T-bond |
| 91 | 2 | X | X | GS10 | 10-Year T-bond | 10 yr T-bond |
| 92 | 2 | | X | AAA | Aaa Corporate Bond Yield | Aaa bond |
| 93 | 2 | | X | BAA | Baa Corporate Bond Yield | Baa bond |
| 94 | 1 | | | COMPAPFF | CP - FFR spread | CP-FF spread |
| 95 | 1 | | | TB3SMFFM | 3 Mo. - FFR spread | 3 mo-FF spread |
| 96 | 1 | | | TB6SMFFM | 6 Mo. - FFR spread | 6 mo-FF spread |
| 97 | 1 | | | T1YFFM | 1 yr. - FFR spread | 1 yr-FF spread |
| 98 | 1 | | | T5YFFM | 5 yr. - FFR spread | 5 yr-FF spread |
| 99 | 1 | X | | T10YFFM | 10 yr. - FFR spread | 10 yr-FF spread |
| 100 | 1 | | | AAAFFM | Aaa - FFR spread | Aaa-FF spread |
| 101 | 1 | | | BAAFFM | Baa - FFR spread | Baa-FF spread |
| 103 | 5 | | X | EXSZUS | Switzerland / U.S. FX Rate | Ex rate: Switz |
| 104 | 5 | | X | EXJPUS | Japan / U.S. FX Rate | Ex rate: Japan |
| 105 | 5 | X | X | EXUSUK | U.S. / U.K. FX Rate | Ex rate: UK |
| 106 | 5 | | X | EXCAUS | Canada / U.S. FX Rate | EX rate: Canada |

Table D.7. Prices

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|-----------------|----------------------------------|------------------------|
| 107 | 5 | X | X | PPIFGS | PPI: Finished Goods | PPI: fin gds |
| 108 | 5 | | X | PPIFCG | PPI: Finished Consumer Goods | PPI: cons gds |
| 109 | 5 | | X | PPIITM | PPI: Intermediate Materials | PPI: int materials |
| 110 | 5 | | X | PPICRM | PPI: Crude Materials | PPI: crude materials |
| 111 | 5 | X | | oilprice | Crude Oil Prices: WTI | Spot market price |
| 112 | 5 | | | PPICMM | PPI: Commodities | PPI: nonferrous |
| 113 | 1 | | | NAPMPRI | ISM Manufacturing: Prices | NAPM com price |
| 114 | 5 | X | X | CPIAUCSL | CPI: All Items | CPI-U: all |
| 115 | 5 | | | CPIAPPSL | CPI: Apparel | CPI-U: apparel |
| 116 | 5 | | | CPITRNSL | CPI: Transportation | CPI-U: transp |
| 117 | 5 | | | CPIMEDSL | CPI: Medical Care | CPI-U: medical |
| 118 | 5 | | | CUSR0000SAC | CPI: Commodities | CPI-U: comm. |
| 119 | 5 | | | CUUR0000SAD | CPI: Durables | CPI-U: dbles |
| 120 | 5 | | | CUSR0000SAS | CPI: Services | CPI-U: services |
| 121 | 5 | | | CPIULFSL | CPI: All Items Less Food | CPI-U: ex food |
| 122 | 5 | | | CUUR0000SA0L2 | CPI: All items less shelter | CPI-U: ex shelter |
| 123 | 5 | | | CUSR0000SA0L5 | CPI: All items less medical care | CPI-U: ex med |
| 124 | 5 | | | PCEPI | PCE: Chain-type Price Index | PCE defl |
| 125 | 5 | | | DDURRG3M086SBEA | PCE: Durable goods | PCE defl: dbes |
| 126 | 5 | | | DNDGRG3M086SBEA | PCE: Nondurable goods | PCE defl: nondble |
| 127 | 5 | | | DSERRG3M086SBEA | PCE: Services | PCE defl: service |

Table D.8. Stock Market

| <i>Series id</i> | <i>Tcode</i> | <i>Medium</i> | <i>Intermediate</i> | <i>FRED</i> | <i>Description</i> | <i>GSI:Description</i> |
|------------------|--------------|---------------|---------------------|---------------|---------------------------|------------------------|
| 81 | 5 | X | X | S&P 500 | S&P: Composite | S&P 500 |
| 82 | 5 | | X | S&P: indust | S&P: Industrials | S&P: indust |
| 83 | 1 | | X | S&P div yield | S&P: Dividend Yield | S&P div yield |
| 84 | 5 | | X | S&P PE ratio | S&P: Price-Earnings Ratio | S&P PE ratio |